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Abstract

In an economy with a seniority wage system, elderly workers are subject to greater income risks when they lose their jobs than young workers are. This paper investigates: (1) whether we can observe the age dependence of idiosyncratic income risks; and (2) the importance of age dependecne for the evolution of inequalities in consumption using Japanese micro data. Our estimation of the income process demonstrates a strong age dependence of income risks; at the age of 48, the variance of permanent income shocks begins to increase, which creates a nonlinear age–variance profile of income. This paper also uses structural estimation of a precautionary savings life cycle model to demonstrate that the nonlinearity in the income process is crucial for understanding the evolution of the consumption inequalities over age.

JEL Classification: D12, D31, D52, E21

Key Words: Income risk, buffer stock savings, consumption inequality, method of simulated moments.

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1 Introduction

It is widely recognized that idiosyncratic labor market uncertainty is one of the most important determinants of household consumption and savings when the market is incomplete. If uninsured income risks contain a permanent component, as Deaton and Paxson (1994) demonstrated, both income and consumption variances grow as people get older. The gap between the two variances, given age, depends on various characteristics of the economy such as household preferences, the degree of market incompleteness, and the magnitude of the income risks. An increasing number of papers regard the gap as a precious source of information for investigating the economy with an incomplete capital market. Blundell and Preston (1998) was one of the first studies that investigated the relationship between the evolutionary patterns of consumption and income variances over age to identify permanent income shocks. With a calibrated model, Storesletten, Telmer and Yaron (2004b) demonstrated the importance of a social security system to replicate the observed pattern of consumption and income inequalities in the United States.

In many papers of calibration or structural estimation of household consumption, the specification of income process is crucial, and the variance of permanent shocks is especially important. Quite naturally, extensive attempts have been conducted to estimate the household income process in detail.[1] Banks, Blundell and Brugiavini (2001) estimated the income process of households in Great Britain using an ARCH structure. Applying PSID, Meghir and Pistaferri (2004) estimated a more general ARCH process. Both studies revealed significant changes in income volatilities over the years. In addition, the latter paper considered the age effects in permanent income shocks and demonstrated that age in the USA does not influence the variance of permanent components in the income process.

The results of Meghir and Pistaferri (2004) suggest that, other things being equal, elderly workers confront the same degree of permanent income shocks as younger workers. Storesletten, Telmer and Yaron (2001) reported that inequalities in hours of work are approximately constant across age, suggesting that it is not necessary to treat elderly workers differently from younger workers in the calibrated model.[2]

[1] Liliard and Weiss (1979), MaCurdy (1982), and Abowd and Card (1989) are particularly influential works in the field. A recent study by Gouvenen (2005) allows for various heterogeneity in estimating the income process.

[2] Except for the difference as a result of different life stages.

Many empirical papers in labor economics propose different stories for the income process. After the seminal work by Topel (1991), many attempts have been conducted to investigate the relationship between the accumulation of firm specific skills and the wage level. Recent careful investigation by Abowd and Kang (2002) demonstrated that a worker with 10 years of tenure is paid about 6.9 percent more than a worker with one-year tenure. If tenure is an important factor for the wage level, the risk of job loss will be more costly for elderly than young workers. Ohtake and Kawaguchi (2005) found that elderly workers in Japan confront higher probabilities of wage cuts and lower probabilities in finding new jobs, which suggests that elderly workers face greater risks in permanent income. If the variance of permanent shocks in income depends on age, the age-variance profile of income and consumption takes a nonlinear form.[3] Deaton and Paxson (1994) and Ohtake and Saito (1998) demonstrated that the age-variance profile of income is linear in the United States, but is not linear in other economies such as Japan and Taiwan.

This paper investigates: (1) whether the idiosyncratic income risks depends on age; and (2) the importance of age dependence for evolution of inequalities in consumption. We take advantage of a large Japanese micro data set that covers the years 1984-1999. As the data set contains variables before and after the bubble economy, we can also analyze the effects of drastic macroeconomic changes to the household income process.[4] Following a procedure similar to that of Storesletten, Telmer and Yaron (2004a), we estimate the income process of Japanese households. The procedure adopted in this paper is different from previous studies in that it allows for the age dependence of permanent income shocks. Our estimation of the income process of Japanese households demonstrated strong age dependence of income risks. As suggested by Ohtake and Saito (1998), at the age of 48, the variance of permanent income shocks begins to increase, which creates a nonlinear age-variance income profile.

Following the results of the estimations of the income process, we investigate whether the age dependence of the permanent income shocks is important for explaining the evolutionary pattern of consumption inequalities in Japan. A structural estimation of

[3]As Deaton and Paxson (1994) demonstrated, the slope of the age-variance profile is the variance of the permanent shocks. If the variance of the permanent component is constant across age, the slope is constant.

[4]While the average growth rate of real GDP during 1984-1990 was 4.7 percent, during the so called "lost decade", 1990-1999, real GDP growth rate was reduced to 1.2 percent.

consumers' dynamic behavior is conducted for this purpose. We build a life cycle saving model with an incomplete capital market that is similar to the buffer stock saving model by Carroll (1997) and Storesletten et al. (2004b). The structural estimation reveals that the model can replicate the changes in variances of consumption over age in Japan, under plausible parameter values only when we consider the age dependence of income shocks. Hence, consideration of nonlinearity in age–variance profile is crucial for understanding the consumption behavior of Japanese households.

Our paper makes a methodological contribution to the literature by documenting inequalities in consumption. We use simulated second moments to conduct statistical inferences of the model. Inferences based on second moments have several advantages over first moments or calibrations without inferences. As Kydland and Prescott (1982) argued, second moments are expected to be less vulnerable to measurement errors than first moments. The level of consumption likely includes age-specific factors that are absent from our model such as health, education, and housing-related expenditures. Although measurement errors have a direct effect on the level of consumption, we expect their impacts on second moments to be less serious than on first moments. In addition, we conduct statistical inferences on several parameter values such as the degree of risk aversions. Household attitudes toward risk are crucial for explaining consumption smoothing and inequalities. By conducting inferences over the parameter, we can introduce "metrics" to evaluate the sensitivity of the estimated value of the risk aversion in our empirical analysis.

The paper is organized as follows. Section 2 discusses our data and briefly explains some characteristics of Japanese households. Section 3 determines the estimates of the income process. Section 4 introduces a dynamic model of consumption with an incomplete capital market. Section 5 presents the procedures of our structural estimation. Section 6 discusses the estimation results, and the final section concludes.

2 Data

The data we employ in this paper are microdata from the National Survey of Family Income and Expenditure (NSFIE), conducted by the Japanese government every five years. The four surveys conducted in 1984, 1989, 1994, and 1999 are used for this project. The surveys contain information on the income earned in the previous year and the three-month average expenditure between September and November of each surveyed year. The total sample size for each survey of the NSFIE is about 50,000, making it one of the largest household surveys in the world with detailed information on income and consumption. The details regarding the survey, such as the sampling procedures and definitions of terms, are described in volumes published for each survey by the Statistics Bureau, Ministry of Public Management, Home Affairs, Posts and Telecommunications. The information is available in both Japanese and English.[5]

We selected a sample of households according to the following criteria: (1) households with two or more members;[6] (2) where the household head is a man aged between 25 and 70;[7] (3) with a household head who is not engaged in the agricultural sector; (4) with a head who is not self-employed or a company director; (5) with fewer than nine family members;[8] and (6) without any missing data. We remove self-employed or company directors from our sample because their incomes reflect corporate performances rather than as being an employee, which substantially increases their income variances. The number of observations we use were 39,030 (1984), 41,558 (1989), 44,077 (1994), and 38,721 (1999).

Our empirical analysis is divided into three steps. First, we create a series of variances of consumption and income for every age and year, which provides us with cohort level data. To control for observable household characteristics such as household composition and location, we regress logged consumption and income on several variables. The variances of the residuals are used for the remaining steps. Second, following Storesletten et al. (2004a), we estimate the income process using the estimated variances from the first step. Finally, we estimate a dynamic model of consumption based on the variances of consumption obtained from the first step and the income process estimated in the second step.

[5]Hayashi, Ando and Ferris (1988) and Hayashi (1995) describe in detail the features of NSFIE that make the survey suitable for analysis of consumption and saving behaviors. Several shortcomings, however, are noticeable. First, the NSFIE does not provide panel data; therefore, we cannot control individual fixed effects directly. Second, the NSFIE does not contain information on educational background. Finally, the data in 1984 do not include information on the size of the firms in which individuals are working.

[6]Single households are surveyed only in September. As other households report the average consumption between September and November, we cannot treat singles and others alike. In this paper, we have removed singles from our sample.

[7]We eliminate observations whose household head is younger than 25 or older than 70 because the number of observations for such younger or older households is small.

[8]There are fewer than 100 households that have more than nine members each year.

To obtain the series of consumption and income variances at cohort level, we regress logged household income and consumption on (1) constant; (2) dummy variables for the number of family members; and (3) prefecture and large city dummies.[9] In a regression of household consumption, we add the number of family members who are younger than 15 to the regressors.[10] Figure 1 reports the variances of residuals obtained from the income regressions. The age-variance profile is upward sloping, which suggests the existence of permanent income shocks. Another noteworthy feature is the shape of the profile. Up to the age of 60, the profile looks convex, that is, the slope is increasing. Since the shapes of the four curves are very similar to each other, the convexity is hardly a by-product of cohort effects.[11] After the age of 60, the variance ceases to increase and remains at a high level. Figure 2 exhibits the age-variance profile of consumption. The convexity of the profile is much clearer in consumption than in income. In all the years surveyed, the consumption variance decreases up to the age of 40, after which it begins to increase rapidly at an accelerating rate.[12]

Because people can receive a public pension after the age of 60, for the remainder of the paper we further restrict our sample to households with the head aged between 25-60, which provides us with $34,991$ (1984), $35,663$ (1989), $36,084$ (1994), and $29,675$ (1999) observations.

3 The Income Process

Following Storesletten et al. (2004a), suppose that an individual income can be decomposed into several components. Denoting the natural logarithms of income for the ith household of age h at time t as y_{it}^h, and observable components such as location and gender as $X_{it}^h \beta_t^h$, and the unobserved components u_{it}^h, the income process admits following

[9]Because we use four NSFIE surveys, the number of regressions is $184 \, (= (70 - 25 + 1) \times 4)$.

[10]When we calculate the logged consumption, we subtract education and medical expenses because the model we use in a later section does not contain medical risks or offspring-related expenditure.

[11]Since the survey is conducted every five years and we have only four surveys, controlling cohort effects as in Deaton and Paxson (1994) by regressing cohort dummies would be difficult in this paper. Rather, to eliminate cohort effects, we take differences of income variances at each cohort in the next section.

[12]This paper is not the first to point out the convexity in both income and age-variance profiles in Japan. Deaton and Paxson (1994) found that households in Taiwan have a similar shape. Ohtake and Saito (1998) suggested that Japanese labor customs such as frequent promotions and layoffs after age of 40 are one cause of this convexity.

decomposition,

$$y_{it}^h = X_{it}^h \beta_t^h + u_{it}^h. \tag{1}$$

When X_{it}^h includes employment status, the residual does not directly reflect the income risks caused by unemployment. In this section, we investigate (1) under several different specifications. A process for u_{it}^h is specified as:

$$u_{it}^h = \alpha_i + \varepsilon_{it} + z_{it}^h, \tag{2}$$

$$z_{it}^h = \rho z_{it-1}^{h-1} + \eta_{it}^h. \tag{3}$$

Where α_i is the time invariable household's fixed effect, ε_{it} is the i.i.d. transitory shock, and z_{it}^h is the permanent component. We assume α_i, ε_{it}, and η_{it}^h are independent of each other, and $\alpha_i \sim N\left(0, \sigma_a^2\right)$, $\varepsilon_{it} \sim N\left(0, \sigma_\varepsilon^2\right)$, $\eta_{it}^h \sim N\left(0, \sigma_{\eta,h,t}^2\right)$. ρ is a scalar that determines the persistency of the income shock through z_{it}^h. We also assume $z_{it}^{\underline{h}} = 0$ where we set $\underline{h} = 24$. That is, at the age of 25, people enter the labor market and begin to be exposed to income risks.

Because the NSFIE does not trace the same households over different surveys, we cannot use the first difference of income within a household to control for household fixed effects as in other works such as Blundell and Preston (1998). We demonstrate that, with some assumptions, ρ and several variances such as $\sigma_{\eta,h,t}^2$ can be identified from the information on the dynamics of variances at cohort level. First, we take the variance of (2) and (3),

$$var\left(u_{it}^h\right) = \sigma_a^2 + \sigma_\varepsilon^2 + var\left(z_{it}^h\right),$$

$$var\left(z_{it}^h\right) = \rho^2 var\left(z_{it-1}^{h-1}\right) + \sigma_{\eta,h,t}^2.$$

Combining the above two equations, we can develop the following expression,

$$var\left(u_{it}^h\right) - \rho^2 var\left(u_{it-1}^{h-1}\right) - \sigma_{\eta,h,t}^2 - \left(1-\rho^2\right)\left(\sigma_a^2 + \sigma_\varepsilon^2\right) = 0.$$

If $\rho = 1$, as in Deaton and Paxson (1994), the above equation becomes simple,

$$var\left(u_{it}^h\right) - var\left(u_{it-1}^{h-1}\right) = \sigma_{\eta,h,t}^2. \tag{4}$$

The equation implies that in each cohort, the variance of income increases by the variance of permanent shock, $\sigma_{\eta,h,t}^2$, each year.

If $\sigma_{\eta,h,t}^2$ is constant over time and age, the income variance grows at a constant rate. Hence, the age-variance profile becomes a straight line. Many empirical analyses such as

Meghir and Pistaferri (2004) demonstrated that the linear profile is a good approximation for the US household.[13] However, as previously discussed, under some labor contract regimes such as the seniority wage and lifetime employment systems, the profile becomes nonlinear. Obviously, Figure1 demonstrates for Japan a convex relationship consistent with findings by Deaton and Paxson (1994) and Ohtake and Saito (1998).

Several procedures capture the nonlinear relationship between age and income variance. The first is to allow ρ to be different from unity. If $\rho > 1$, the age-variance profile has an increasing slope. The shape in Figure 2, however, suggests a more complicated nonlinear income process in Japan because the nonlinearity in consumption variances is greater than that of income. Ohtake and Saito (1998) provided several possible explanations for the nonlinearity, for example, a special promotion system within a firm. Although many empirical and theoretical investigations have been conducted on Japanese employment customs, there is no consensus on the mechanism that characterizes the Japanese labor market.[14] Thus, we have decided that rather than depending on specific behavioral models in labor markets, we attempt to model the data described by Ohtake and Saito (1998). Consequently, as the second model, we assume $\rho = 1$ and the constant variance, $\sigma^2_{\eta,h,t} = \sigma^2_\eta$, until a certain "Turning Age" and allow for growth in variances, $\sigma^2_{\eta,h,t}$, after this age.

We estimate the following two models.

[Model1]

$$var\left(u^h_{it}\right) - var\left(u^{h-1}_{it-1}\right) - \sigma^2_{\eta,h,t} = 0,$$

$$\sigma^2_{\eta,h,t} = \sigma^2_\eta + \delta_1 1994dummy + \delta_2 1999dummy \text{ if } h \leq \text{ Turning Age},$$
$$\sigma^2_{\eta,h,t} = \left(\sigma^2_\eta + \delta_1 1994dummy + \delta_2 1999dummy\ \right)^{(h-TurningAge)g} \text{ if } h > \text{Turning Age},$$

[Model2]

$$var\left(u^h_{it}\right) - \rho^2 var\left(u^{h-1}_{it-1}\right) - \sigma^2_{\eta,h,t} - \left(1-\rho^2\right)\left(\sigma^2_a + \sigma^2_\varepsilon\right) = 0,$$

$$\sigma^2_{\eta,h,t} = \sigma^2_\eta + \delta_1 1994dummy + \delta_2 1999dummy\ .$$

We use the maximum likelihood method to estimate parameters, $\left(\sigma^2_\eta, \delta_1, \delta_2, g, \text{Turning Age}\right)$ for Model 1, and $\left(\sigma^2_\eta, \delta_1, \delta_2, g, \text{Turning Age}\right)$ for Model 2. In Model 2, we need information of $\sigma^2_\alpha + \sigma^2_\varepsilon$ to conduct our estimation. Because the NSFIE is cross–sectional, we

[13] Storesletten, Telmer and Yarons' (2004a,b) estimation of ρ with US data is 0.9989.

[14] Hashimoto and Raisian (1985) and Aoki (1988) are very influential works on this issue.

cannot identify σ_α^2 or σ_ε^2, directly. We can estimate the variance of permanent shock, however, as long as the variances of fixed effects, σ_α^2, and temporary shocks, σ_ε^2, are constant over time. Assuming $z_{it}^{24} = 0$, it is possible to derive $\sigma_\alpha^2 + \sigma_\varepsilon^2 = var\left(u_{it}^{25}\right)$, which enables us to estimate Model 2 from our data set. Dummies representing years are included to capture the year effects on permanent income shocks.[15] 1994$dummy$ and 1999$dummy$ take unity if the data are in the 1994 and 1999 surveys, otherwise they are zero.[16] As the variance, $var\left(u_{it}^h\right)$, we use the variances of the residuals of (1) for each surveyed year and age. We construct eight different types of residuals by using different controls for (1).[17]

Tables 1 and 2 report the estimation results for Model 1 and Model 2, respectively. Except for Spec 1, we can reach an interior optimum for the Turning Age at around age 48. The growth rate of the variance after the turning age us significantly positive for all the specifications. The significance of the growth rate, g, implies that we can reject the hypothesis that the income process is linear, such as $\sigma_{\eta,h,t}^2 = constant$ for all ages. The year effects for 1999 are extremely small and insignificant for most cases. We can observe an increase in permanent shocks in income only in Spec 7 and Spec 8. This implies we can observe an increase in volatility of permanent income shock only within the same industry groups, or the same types of jobs. Meanwhile, 1994 effects are generally positive and significant except for Spec 3.

According to Table 2, the AR1 coefficient of the lagged variance, ρ, exceeds unity. This is a natural consequence from the convex age-variance profile. The 1994-1999 dummy is significantly positive only for Specs 7 and 8. Considering the results in Tables 1 and 2 together, we conclude that an increase in inequality in permanent income arises only within industries or job types. To evaluate the relative explanatory power of Models 1 and 2, we conduct Vuong's test for non-nested models. The result is reported in Table 3. As we expect from the differences in likelihood values, Model 1 is always selected.

Spec 4 to Spec 8 control for employment status of the household head, which implies that the variances obtained from these specifications do not reflect income risks due to

[15] Although we need parameter values for annual data for later sections, the NSFIE is available only with five year intervals. In the estimation, we use $var\left(u_{it}^h\right) - var\left(u_{it-5}^{h-5}\right) - 5\sigma_{\eta,h,t}^2 = 0$ for Model 1, and $var\left(u_{it}^h\right) - \rho^{10} var\left(u_{it-5}^{h-5}\right) - \sigma_{\eta,h,t}^2 - \left(1 - \left(\rho^8 + \rho^6 + \rho^4 + \rho^2\right)\right)\left(\sigma_a^2 + \sigma_\varepsilon^2\right) = 0$, for Model 2.

[16] Because we utilize the information on the changes in variances in the same cohort during the five year interval between the surveys, 1994$dummy$ and 1999$dummy$ capture the changes in variances between 1989-1994, and 1994-1999.

[17] Appendix Table 1 contains a detailed explanation for each specification.

employment status. To capture the income risks from the labor markets, we use Spec 1 to Spec 3. Spec 3 controls for family composition and location of the households. Although these are not completely exogenous, they are not likely to be the main cause of the income risks. We consider them as a part of the family fixed effects. Therefore, hereafter, we use the results from Spec 3, in which the year effects are not significant. Table 4 reports the results of Model 1 estimation without the year effects. There are no substantial differences between Tables 1 and 4 in the Turning Age or g.

4 The Household Model

In this section, to explain the nonlinearity of the variance of the age log-consumption profiles, we introduce a consumption/saving model developed by Carroll (1997) and extended by Gourinchas and Parker (2002) and Cagetti (2003). In the model, because of missing insurance markets, agents confront idiosyncratic earnings risks and accumulate savings as a buffer stock for precautionary motives when they are workers.[18] If a negative income shock is realized, the agent deccumulates his/her assets for the purpose of smoothing the consumption path over the life cycle. This is why the model is known as the "buffer-stock saving model." Carroll (2004) carefully explains the theoretical properties of the model. One of our objectives in this paper is to investigate whether the model can replicate the consumption inequalities measured by the variances of logarithms of consumption in Japan under plausible parameters.

Consider a household's lifecycle consumption/saving problem. Although households live for at most T periods, they face survival probabilities $\{s_h\}_{h=1}^{T}$ for each age h. Consequently, death is uncertain and a household may die with an accidental bequest. The household supplies labor with an inelastic supply curve for $R < T$ and obtains labor earnings, and retires after R years. The household maximizes its expected utility:

$$U\left(\{C_h\}_{h=1}^{T}\right) = E\sum_{h=1}^{T}\beta^{h-1}\frac{C_h^{1-\gamma}}{1-\gamma}S_h, \text{ where } S_h = \prod_{j=1}^{h}s_j,$$

where $\beta > 0$ is a discount factor and S_h is a cumulative death probability for age h. We assume that the instantaneous utility function is of the constant relative risk aversion form, i.e., ($u(C_h) = \frac{C_h^{1-\gamma}}{1-\gamma}$). At the beginning of age h, each household has some cash

[18] For empirical studies on the market incompleteness in the United States, see Cochrane (1990) and Mace (1991); in the case of Japan, see Kohara, Ohtake, and Saito (2002). Storesletten et al. (2001) investigated the market incompleteness from a different point of view.

on hand (net worth plus current income), X_h. We also assume they have some financial wealth when they enter the economy. The cash on hand is, then, allocated between consumption C_h and savings:

$$C_h = X_h - W_h,$$

where W_h represents the financial wealth at the end of age h. The wealth yields interest in the next period, and the gross interest rate is given by $(1 + r)$. We assume that the interest rate is constant over time and for each cohort. We also assume all agents face liquidity constraints, (i.e., $W_h \geq 0$).[19] The next period's cash on hand is given by:

$$X_{h+1} = (1 + r)W_h + Y_{h+1},$$

where, during working age, the household receives labor earnings, Y_{h+1}. As already mentioned, households face idiosyncratic labor income risks, which can be decomposed into permanent shocks, ψ_{t+1}, and transitory components of the shocks, ξ_{h+1}. The permanent income component evolves according to a deterministic growth rate G_h and a permanent shock, ϕ_h. As shown in the previous section, as the variances of consumption in Japan are convex over age, we have to consider age-specific permanent income shocks; this differentiates our model from previous research such as Gourinchas and Parker (2002) and Cagetti (2003).

The average income growth rates can be easily obtained by the NSFIE. We assume that the transitory and permanent shocks follow the log-normal distribution with means $(-\frac{\sigma^2_{\xi_{h,t}}}{2}, -\frac{\sigma^2_{\phi_{h,t}}}{2})$ and variances $(\sigma^2_{\xi_t}, \sigma^2_{\phi_{h,t}})$. By increasing the number of simulations, the average income in the model approaches the mean of the data-generating process. Note that the means and variances depend upon not only age h but also year t.

$$Y_{h+1} = \psi_{h+1}\xi_{h+1}, \text{ if } h + 1 \leq R \tag{5}$$

$$\psi_{h+1} = G_{h+1}\phi_{h+1}\psi_h, \tag{6}$$

$$\ln \phi_h \sim \mathcal{N}\left(-\frac{\sigma^2_{\phi_{h,t}}}{2}, \sigma^2_{\phi_{h,t}}\right), \ \ln \xi_h \sim \mathcal{N}\left(-\frac{\sigma^2_{\xi_t}}{2}, \sigma^2_{\xi_t}\right),$$

where initial permanent component ψ_1 is determined by the "fixed effect," α, which cannot be shared by precautionary savings.

A retiree draws a social security payment from the government. The amount of the pension is determined by the permanent income level at the time of retirement multiplied

[19] If there is a non-zero probability that an agent's income becomes zero, the no-Ponzi game condition prohibits the household from borrowing.

by b_t:

$$Y_{h+1} = b_t \psi_{h+1}, \text{ if } h+1 \geq R+1 \tag{7}$$

where b is determined by the replacement rate observed in Japan.

Defining the value function of our problem as $V_h(X_h, \psi_h)$, where the cash on hand, X_h, and the permanent income, ψ_h, are state variables, we define a household's optimization problem as a dynamic programming of the following form:

$$V_h(X_h, \psi_h) = \max_{C_h} \left\{ \frac{C_h^{1-\gamma}}{1-\gamma} + s_h \beta E_h V_h(X_{h+1}, \psi_{h+1}) \right\} \tag{8}$$

$$\text{subject to}$$

$$X_{h+1} = (1+r)[X_h - C_h] + Y_{h+1}, \text{ (5), (6) and (7).}$$

To minimize the number of state variables, by using homogeneity of the value function, we divide both sides of the Bellman equation by ψ_h. Then, the model becomes a simple form such as [20]:

$$v_h(x_h) = \max \left\{ \frac{c_h^{1-\gamma}}{1-\gamma} + s_h \beta E_h \Gamma_{h+1} v_h(x_{h+1}) \right\} \tag{9}$$

$$\text{subject to}$$

$$x_{h+1} = \frac{(1+r)}{\Gamma_{h+1}}[x_h - c_h] + \xi_{h+1}, \text{ if } h \leq R, \tag{10}$$

$$x_{h+1} = \frac{(1+r)}{\Gamma_{h+1}}[x_h - c_h] + b, \text{ if } h > R, \tag{11}$$

where $\Gamma_{h+1} \equiv G_{h+1}\phi_{h+1}$, and the lower cases are the normalized variables (e.g., $c_h \equiv \frac{C_h}{\psi_h}$, $x_h \equiv \frac{X_h}{\psi_h}$). The corresponding policy functions are denoted by $c_h = c_h(x_h)$. The Euler equation is, then, given by:

$$\begin{aligned} c_t^{-\gamma} &\geq s_h \beta \Gamma_{h+1}^{-\gamma}(1+r) E_h c_{h+1}^{-\gamma} \\ &= s_h \beta \Gamma_{h+1}^{1-\gamma} \frac{(1+r)}{\Gamma_{h+1}} E_h \left[c_{h+1} \left(\frac{(1+r)}{\Gamma_{t+1}}[x_h - c_h] + \xi_{h+1} \right)^{-\gamma} \right] \end{aligned} \tag{12}$$

It is well known that the problem has no analytically tractable solution. Therefore, we need to resort to a numerical method. We apply "the endogenous gridpoints method" to compute the policy function from the Euler equation (12). For more detailed procedures, see the appendix or Carroll (2005). Figure 3 depicts the normalized consumption functions (the policy functions, $c(x_h)$) for several ages. As Carroll and Kimball (1996)

[20] For normalization of the model, see the appendix.

demonstrate, the function is continuous and concave. The horizontal axis represents the normalized cash on hand, x_h, for each age, and the corresponding vertical figures are the consumption. Because of lifecycle motives, young consumers have high propensity to consume, even if the normalized cash on hand is the same. The shape of concavity is derived from the formulation of the income process and the liquidity constraint.

5 The Method of Simulated Moment

Following from the previous section, given the estimated income process obtained in Section 3, we examine how well the model can explain the variances profile of age-consumption by using the simulated moments method. Note that our objective in this paper is *not* to estimate exactly the fundamental parameters of the Japanese economy, but to investigate the consistent parameters with regard to income and consumption inequalities. In other words, we investigate the differences between the consumption and income inequalities by using the buffer stock saving model under plausible parameters. In the next section, we also consider whether there are changes in the relative risk aversion among data years in the NSFIE.

Our estimation procedures follow previous research such as Gourinchas and Parker (2002) and Cagetti (2003). The method employed is named "the method of simulated moments (hereafter, MSM)" and was employed in the aforementioned studies by proceeding with the two following steps. First, given a fundamental parameter, we solve the model numerically and generate several simulated data. Second, the studies seek to find the parameters that match the simulated data with those of the real economy in the United States. By using the method of simulated moments, the studies estimate the fundamental parameters (β, γ) by matching the averages of age-consumption or age-wealth profiles, which implies that Gourinchas and Parker (2002) and Cagetti (2003) used the first moments of the micro data. In our estimations we use the second moments of the cross-sectional consumption distribution (i.e., the variances) instead of the first moments. We estimate the fundamental parameter of our model, γ, using the second-order moments.

Since it is practically impossible to solely estimate *all* parameters of the economy using the method of simulated moments, we conducted both estimation and calibration for our fundamental parameters. Following Gourinchas and Parker (2002), we take two steps for the estimation. As a first-stage estimation, to solve the model numerically, we calibrate and estimate some important parameters from the NSFIE. As a second step, we estimate the relative risk aversion by using those parameters.

5.1 First Stage Estimation

For computing the policy function of the household's problem, we have to specify several parameters such as the growth rate of average income $\{G_{h,t}\}_{h=1}^{T}$, the stochastic income process $(\sigma^2_{\xi_t}, \sigma^2_{\phi_{h,t}})$, the average and variances of initial financial wealth $(\bar{x}_{25,t}, \sigma^2_{\bar{x}_{25,t}})$, the discount factor β, the survival probabilities $\{s_h\}_{h=1}^{T}$, the replacement rate b_t, and the fraction of the fixed effect σ^2_{α}.

a. Growth rate of income: Computing the growth rates $\{G_{h,t}\}_{h=1}^{T}$ involves the average age-income profiles (i.e., $G_{h+1,t} = \bar{Y}_{h+1,t}/\bar{Y}_{h,t}$) for each year. For each age and each year, we have computed the average income, $\bar{Y}_{h,t}$, by controlling the same variables as in the previous section. Afterwards, using fourth order polynomials, we smooth out the age-income data and obtain the profiles as in Figure 4.

b. Stochastic income process: The idiosyncratic income process in the model is specified entirely the same as in Section 3. Our formulation of the stochastic income process can be easily transformed into the equation (1), (2), and (3) defined in Section 3. In the MSM estimation, we use estimated results of only the case of $\rho = 1$, because of "the curse of dimensionality."[21] That is, we need to solve a two-dimensional Bellman's equation even after the normalization if ρ is not equal to unity, which is costly to solve numerically. We use Spec 3 without the year dummy represented in Table 4. Following that, we can transform the idiosyncratic income risk process into the model in Section 3 such that:

$$\ln Y_{h+1} = \ln \psi_{h+1} + \ln \xi_{h+1},$$
$$\ln \psi_{h+1} = \ln G_{h+1} + \ln \phi_{h+1} + \ln \psi_h,$$
$$\mathrm{var}(\alpha^h_{i,t}) + \mathrm{var}(\varepsilon^h_{i,t}) = \mathrm{var}(\ln \xi_{t,h}),\ \ z^h_{i,t} = \mathrm{var}(\ln \psi_h),\ \ \eta^h_{i,t} = \mathrm{var}(\ln \phi_{t,h})$$

c. Initial financial wealth: Households hold some savings when they enter the economy. We assume that all agents start their economic activity at 25 years old and each period in the model is set to be one year. Because financial wealth at $h = 1$ must be determined outside the model, we have to estimate and calibrate the distribution of

[21] For approximation of the distribution of ϕ and ξ, we use the Gauss-Hermite quadrature method, and take 32 and 40 gridpoints respectively. For details of the quadrature method, see Judd (1998)

the wealth for a 25-year-old, i.e., $x_{1,t} = X_{1,t}/\psi_{1,t}$. It is not easy, however, to directly associate the wealth derived from our model with that observed from the data. We have introduced a liquidity constraint in our model, which implies that the financial wealth must be non-negative. In the NSFIE, there are many households with negative assets; one possible way to eliminate them is to remove them. Such a procedure, however, overestimates wealth at 25 years old. Therefore, first, we estimate averages and variances of the normalized cash on hand with negative assets, as in Table 5. Second, in the simulation process, agents with negative assets are replaced with zero assets.

d. Discount factor: Note that the discount factor β is a fundamental parameter, and is estimated in other studies. We do not estimate β and γ simultaneously because, as increasing both variables changes the consumption variances in the same direction (the variances decrease), it is difficult to correctly estimate both parameters. Moreover, the discount factor is closely related to the interest rate. We set β as equal to 0.95 and tested the model with several interest rates, $r = \{3\%, 4\%, 5\%\}$. From a theoretical point of view, the interest rate must be lower than the subjective discount rate if agents live forever, even though it is not well known if agents live a finite period.

e. Survival probabilities: The maximum living period is set to be $T = 61$, which implies that an agent lives at most 85 years. Following the Japanese tradition, the retirement age is exogenously determined to be 60 years old (i.e., $R = 36$). The survival rates are taken from the National Institute of Population and Social Security Research (2002).

f. Replacement rate: Retirees finance their expenditure with savings and the public pension. The parameter b is set so that the replacement rate of social security is 59.3%, i.e., $b_t \in \{0.54, 0.59, 0.59, 0.54\}$ for each year. Such values makes the replacement rate (average income after retirement/average income over life cycle) the model case of the Ministry of Health, Labour and Welfare in Japan, namely 59.3%.

g. Fixed effect: It is impossible to distinguish the fraction of the fixed effect and the transitory shocks because of data limitation.[22] Thus, to investigate quantitative consequences of the fixed effects, we estimate the model parameters under different values for the fraction of the fixed effects, $\{0.4, 0.5, 0.6\}$. The fixed effects plus transitory shocks

[22] Recall that the NSFIE is constituted as a repeated cross section: there is no information about the same households' income over several years. Moreover, the NSFIE does not contain data on educational background.

for each year are calculated from the averages of the $25-29$-years-olds.[23]

5.2 Second Stage Estimation: Details of the MSM Procedures

We use the method of simulated moments developed by Pakes and Pollard (1989), Duffie and Singleton (1993), and applied by Gourinchas and Parker (2002) and Laibson, Repetto, and Tabacman (2004), who estimated the consumption function. Suppose that a variance parameter of the log-likelihood function is χ and the relative risk aversion is γ. We search for the parameters that maximize the following log-likelihood function:

$$\mathcal{L} = -\frac{R}{2}\ln(2\pi) - \frac{R}{2}\ln\chi^2 - \frac{1}{2}\sum_t\sum_{h=1}^{R}\left(\frac{1}{\chi^2}g_{t,h}(\gamma)^2\right),\ t \in \{1999, 1994, 1989, 1984\},$$

$$g_{t,h}(\gamma) = \text{var}(\hat{C}_{t,h}) - \text{var}(\tilde{C}_{t,h}(\gamma)), \tag{13}$$

or

$$g_{t,h}(\gamma) = \left[\text{var}(\hat{C}_{t+5,h+5}) - \text{var}(\hat{C}_{t,h}(\gamma))\right] - \left[\text{var}(\tilde{C}_{t+5,h+5}) - \text{var}(\tilde{C}_{t,h}(\gamma))\right], \tag{14}$$

where $\text{var}(\hat{C}_{t,h})$ is the variance of the logarithms of consumption for the Japanese economy in year t at h years old, and $\text{var}(\tilde{C}_{t,h}(\gamma))$ is the variance of logarithms of simulated consumption given γ.[24] Basically, the consumption data are controlled by the same variables as in the previous section plus the number of children.[25]

To compute the likelihood function, we first compute the approximated policy functions for each age based upon the Euler equation (12). Following that, by generating idiosyncratic income processes, we simulate for many agents the variances of age-consumption profiles. The number of agents simulated is set to be 10,000. From the simulated data set, we can easily obtain the variances of logarithms of consumption and compare them to the consumption data.

The basic idea behind the estimation is the same as in Gourinchas and Parker (2002). Meanwhile, because we use the variances of consumption of cross–section distribution, there are two possible ways of estimation. The first approach is, as in the previous section, by using slopes from within the cohort variances as in (14). The other method is using the level information of the logged consumption variances, (13).

[23]On the effect of education for lifetime earnings, see Keane and Wolpin (1997).

[24]Education and medical expenditures are excluded from household consumption in our definition of consumption.

[25]The variances of income and consumption used are placed in Appendix Table 2.

The variances of the estimated parameters are obtained from the inverted Hessian such as:

$$\{\mathbf{I}(\mathbf{\Theta})\}^{-1}\,,\ \{\mathbf{I}(\mathbf{\Theta})\} = \left\{-E\left[\frac{\partial^2 \ln \mathcal{L}(\mathbf{\Theta})}{\partial\mathbf{\Theta}\partial\mathbf{\Theta}'}\right]\right\},$$

where $\Theta = (\gamma, \chi)$ is the vector of structural parameters.

¿From those procedures, we estimate the fundamental parameter γ, and find that the model with nonlinear permanent shocks can explain Japanese consumption inequality very well. We also demonstrate that there is a significant difference between four years' pooled estimation and three years' estimation in which the 1999 year is excluded.

6 Results and Discussion

The objectives for conducting the estimations are not to search for the true values of the parameters, but rather to investigate whether the estimated parameters give us the best fit with the consumption inequality within the plausible range. If our estimated parameters remain within the reasonable range, and if the overall fit is acceptable, we can determine that our structural model can explain the spread of the consumption inequality over the life cycle in Japan.

We establish that the parameters obtained in the MSM are actually plausible. The simulated consumption profile generated by the buffer stock saving model with plausible parameters can accurately explain the observed consumption inequalities over the life cycle only when we take into account the nonlinearity of the stochastic income process. In other words, if the agents confront the constant permanent income risks over age measured by the variances of income, then the generated consumption path does not account for the consumption profile with the plausible parameters.

6.1 Estimation Using the Variances Level of Consumption for Each Cohort

As a starting point, we discuss our estimated results by considering the estimation using the variance level information of each cohort, i.e. the estimation based upon equation (13). This procedure uses all four years variances of age-consumption from 1984 to 1999, and compares them with the data. Each household in 1984, 1989, 1994, and 1999, confronts different earnings profiles, idiosyncratic transitory shocks and fixed effects, and financial wealth at 25 years old. Therefore, we generate four simulated consumption paths and obtain four variance profiles of age–consumption.

As in Figure 5, our simulated results can explain well the variances of the Japanese consumption profile.[26] When young, the variances of consumption seem almost flat around 0.10, which is very close to the corresponding variances of income. At middle age, the consumption inequality sharply rises, and at retirement age, it appears to remain constant. Those characteristics are accurately replicated by the consumption path generated by the policy function and simulation with plausible parameters.

Table 6 reports the estimated results using four years of pooled data and the levels of the variances of age-consumption. The fraction of the fixed effect is determined from 0.4 to 0.6 and the interest rate is set as $r \in \{0.03, 0.04, 0.05\}$. Moreover, we estimate the fraction of the fixed effect by including the fixed effect parameter in the simulated process. It is apparent that the estimated relative risk aversion is very sensitive to the fixed effects and the interest rate, and it takes values from 0.12 to 2.83. Although the estimated risk aversion is within a wide range, we can positively say that those values are plausible and acceptable as demonstrated in the literature. Even though the fixed effects cannot be estimated directly from the NSFIE, the values estimated by the MSM appear to be plausible (about 0.6).[27] If the fraction of the fixed effect is high, the amount of risks households can hedge is small.

What makes the estimated values so different? The answer is found in the amount of savings, which is accumulated for hedging the income risks. As risk aversion increases, households are inclined to hold greater savings. If the interest rate is high, households are willing to save more at the same risk aversion. Therefore, given the relationship between the variances of consumption and income, as the interest rate increases, the risk aversion is estimated to be low, as revealed in Table 6. This suggests that if the relationship between the variances of consumption and income are close to each other (and the discount rates are constant over the four years), then the recent low interest rate in Japan implies that the relative risk aversion might be high.

One of our findings is that when we exclude the variances of consumption in the year 1999 from our estimation, as in Table 7, the estimated risk aversions become higher relative to Table 6. Why does our estimation indicate a higher risk aversion if we omit the consumption data in 1999? Currently, there are many risks in Japan such as employment

[26]For depicting Figure 5, we set the interest rate to be 4 percent and the fraction of the fixed effect as a benchmark is 0.5.

[27]Note that the standard deviation of the risk aversion increases when we include the fraction of the fixed effect in the estimation.

risks and a career crisis related to job change.[28] In our model, low risk aversion implies low savings. In particular, if the agents hold financial wealth poorly, they are unable to share their idiosyncratic income risks. Thus, agents behave in a hand-to-mouth manner, and the variances of income are close to those of consumption. Our result implies that the cash on hand is slightly low at a young age in the year 1999. This results in the three-year estimation of the risk aversion being higher than the four-year estimation.

These findings are supported by Figures 6*a*–6*c*, which depict the financial wealth normalized by annual income as defined in the model for each year and age. As demonstrated, the average net financial asset over income in 1999 is relatively low for the young generation. Thus, the comparatively high estimation of the relative risk aversion in the three-year estimation can account for the low liquidity at a young age in 1999. On the contrary, variances of the normalized net financial assets are high in 1999. For a definition of net financial assets, a housing loan is offset by holding the house itself on the balance sheet. If the housing loan is included in the net financial assets, then the average normalized net financial assets appear to be the same for all years. This implies that housing loans damaged household liquidity.

6.2 Estimation Using Cross-Section Consumption Inequality for Each Year

If the relationship between the income risks and the consumption inequality have ergodicity, the generated cross-section relationships must be the same. As revealed in Figures 1 and 2, it does have ergodicity. Thus, we estimated the relative risk aversion by the MSM using purely cross-section data, and obtained the results presented in Table 8. We used only each year's earnings profile, the initial wealth, the fixed effect, and the transitory shock, and fitted our simulated variances of consumption with the cross-section consumption inequalities for each year. Therefore, we used only cross-section information.

The estimated relative risk aversions are apparently different for each year. In particular, the relative risk aversions are estimated to be high in the 1980s, and low in the 1990s. If the discount rates and the interest rates are the same for each year, the risk aversion is estimated to be high implying that the liquidity to share the risk is sufficient. The interest rates in Japan in the 1980s were high compared to those in the 1990s be-

[28]The 1990s recession of the Japanese economy is known as "the lost decade." For a detailed discussion on the recession, see Hayashi and Prescott (2002).

cause of the financial market bubble. The result is that the true relative risk aversions were higher in the 1990s.

6.3 Estimation Using Within Cohort Effect

One possible objection about the estimation is the use of the “level” of variances as opposed to the slopes of the variances, as in Section 3. By using slopes of the variances of consumption, we can focus on the pure effects of each cohort. Because controlling the variable results in a large specification error for the level of the variances, it is preferable to use the slopes of variances. Therefore, we estimated the risk aversion based upon Equation (14). Although the risk aversions in Table 9 are relatively high compared to those in Table 6, the qualitative properties of the estimation are the same as those given above Figure 7 demonstrates that the simulated consumption variances for the four years replicate the data correctly for each cohort.

6.4 The Importance of Nonlinearity for Japan

Research by Storesletten et al. (2004b) quite successfully explains the US data by adopting the precautionary savings model and linear income process estimated by the US Panel Study of Income Dynamics. However, as demonstrated in the previous section, the variances of Japanese or Taiwanese consumption grow increasingly in middle age. In the buffer stock saving model with linear permanent shock process, as in Figure 8, the corresponding variances profile of age-consumption becomes a concave function over the age.[29,30] Therefore, Storesletten et al. (2004b)'s estimation is likely to perform poorly if we apply the same estimation to the Japanese economy. Although they explain the shape of the age-consumption inequality in the United States, they fail to explain the “level” of variances of logarithms of consumption; as a result Storesletten et al. (2004b) concluded that the social security system is essential for understanding the level of fit.

In Storesletten et al. (2004b), the relative risk aversion is fixed to be 2. Our approach exhibits a sharp contrast. Using a partial equilibrium model, we calibrated the replacement rate from the public pension system in Japan, and estimated the relative risk aversion that is consistent with the Japanese data's shape and level. Using the

[29]Since it is easy to compute the linear regression parameters, they are omitted. Notice that the fixed effects plus transitory shocks have also changed.

[30]For detailed discussion of the concavity of the variances profile of consumption, see Storesletten et al. (2004b).

nonlinear stochastic income process, we are able to replicate the consumption inequality with a plausible risk aversion. Meanwhile, if we assume the variances of permanent shocks to be constant over age, the risk aversion that maximizes the log-likelihood has a value that is unacceptably large (Table 10). Moreover, at least for the Japanese economy, Vuong's test reveals that the nonlinear income process model is superior to the model with linear regression.

7 Concluding Remarks

In this paper, we have investigated the determinants of consumption inequalities in Japan between 1984 and 1999. There are two major findings: (1) Japanese elderly households confront greater permanent income risks than do younger households; and (2) the age dependence of the income risks is crucial for understanding the consumption inequalities in Japan. These findings suggest that when we consider Japanese household behavior toward income risks, we need to take into account labor market customs such as lifetime employment and the seniority wage system.

In addition, we have studied whether the permanent income shock is increasing. Our results demonstrate that, in general, we cannot confirm significant changes in income risks during the sample period, 1984–1999. Significant increases can be observed only within the same industry or the same type of job.

There are several remaining issues to be examined. First, although income inequalities are basically constant over the sample periods for the same age groups, inequalities in consumption are rising particularly among the young generation. We suspect that this is due to an increase in amounts of housing loans. Because of the Bank of Japan's zero interest rate policy and several special tax provisions for housing loans, the amount of debt among young households has increased dramatically, which has likely made young households more vulnerable to income risks and has increased inequalities in consumption even though the income volatilities are unchanged. In addition, other risks such as health risks have to be incorporated into our model. As the number of elderly people increases due to the rapid aging of the Japanese population, the importance of risks related to health and illness is likely to increase for the whole economy. Finally, a more detailed consideration of the labor market is necessary. In this paper, we did not use a specific model for describing age dependence of income risks. Although the difficulty in computation is formidable, we have to incorporate this aspect in detail to understand the household behavior toward income risks.

A Technical Appendix:

In the appendix, we explain the numerical procedures used in this paper. We first demonstrate how the two-dimensional dynamic programming (8) can be normalized to a one-dimensional dynamic programming problem, which is easier to solve numerically. From Equation (5) and (6), by dividing the permanent income ψ_h, we obtain the normalized budget constraint for each h:

$$\begin{aligned} \frac{X_h}{\psi_h} &= (1+r)\left(\frac{X_h}{\psi_h} - \frac{C_h}{\psi_h}\right) + \frac{Y_{h+1}}{\psi_h}, \\ G_{h+1}\phi_{h+1}x_{h+1} &= (1+r)\left(x_h - c_h\right) + G_{h+1}\phi_{h+1}\xi_{h+1}, \\ x_{h+1} &= \frac{(1+r)}{G_{h+1}\phi_{h+1}}\left(x_h - c_h\right) + \xi_{h+1}. \end{aligned}$$

Moreover, by backward induction, we can prove that the value function $V_h\left(X_h, \psi_h\right)$ is a homogeneous function, i.e., $V_h\left(X_h, \psi_h\right) = \psi_h^{1-\gamma} v_h\left(x_h\right)$. Therefore, we get a consumer's problem as in (9), (10), and (11).

Let us define the second term of the right hand side of the Bellman equation (9) as a function of normalized financial wealth w_h:

$$\begin{aligned} \Omega\left(w_h\right) &= \Omega\left(x_h - c_h\right), \\ &= s_h\beta\left[E_h\Gamma_{h+1}v_{h+1}\left(\frac{(1+r)}{\Gamma_{h+1}}w_h + \xi_{h+1}\right)\right]. \end{aligned}$$

¿¿From the envelope condition, the first order condition is to obtain:

$$\begin{aligned} u'\left(c_h\right) &= \Omega'\left(x_h - c_h\right), \\ c_t^{-\gamma} &= s_h\beta\left[E_h\Gamma_{h+1}^{1-\gamma}\frac{(1+r)}{\Gamma_{h+1}}v_h'\left(\frac{R}{\Gamma_{h+1}}w_h + \xi_{h+1}\right)\right], \\ &= s_h\beta(1+r)\left[E_h\left(\Gamma_{h+1}c_{h+1}\right)^{-\gamma}\right]. \end{aligned}$$

Because our consumption problem has a finite horizon, we determine the next period's consumption function by backward induction.

For practical computation, we adopt the endogenous gridpoints method, which is rapid, accurate, and relatively easy to implement. We take gridpoints on a net financial asset $\left\{a_h^1, \ldots, a_h^J\right\}$. We took $J = 200$. Then, we can obtain consumption for each grid such that $c_h^j = \left[\Omega'(a_h^j)\right]^{-\frac{1}{\gamma}}$ for each grid j. Note that the function $\Omega'(a_h^j)$ is computed easily if we approximate the integration of the expectation operator using the Gaussian quadrature method. As $x_h^j = c_h^j + a_h^j$, we have the approximate consumption function $c_h = c_h(x_h)$. For a more detailed procedure, see Carroll (2005).

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Table 1
Estimation of Permanent Income Shocks

	Spec1	Spec2	Spec3	Spec4	Spec5	Spec6	Spec7	Spec8
Variances of Permanent Shocks	0.0000703***	0.00116**	0.0016793***	0.0013497***	0.0017784***	0.0011619***	0.000241	0.00019964***
	(11.13)	(2.17)	(8.09)	(9.32)	(7.25)	(2.34)	(1.31)	(2.72)
Variance Growth Rate after Turning Age	0.21792***	0.23427***	0.2081***	0.2001***	0.23327***	0.2377***	0.40597***	0.38807***
	(154.91)	(5.48)	(21.33)	(66.63)	(23.15)	(5.01)	(4.30)	(9.71)
89--94 year dummy	0.000781***	0.002294**	0.001202	0.0014409*	0.0017396**	0.0015628*	0.0013949**	0.0016127***
	(2.35)	(2.12)	(1.45)	(1.82)	(2.08)	(1.84)	(1.99)	(2.47)
94--99 year dummy	0.000945	0.000072	0.000116	0.001001	0.000810	0.001209	0.0023081***	0.0023298***
	(0.994)	(0.07)	(0.14)	(1.28)	(0.99)	(1.39)	(3.27)	(3.56)
Turning Age	31#	47	48	48	51	49	49	48
σ	0.025004	0.023427	0.021287	0.019761	0.020000	0.018318	0.017384	0.017178
log-likelihood	211.09	215.1038	226.06	232.98	231.86	240.02	244.89	246.00
Fixed Effects & Transitivity Shocks	0.130568	0.119139375	0.100830625	0.100830625	0.0838144	0.07842955	0.07807105	0.0732457

Data : National Survey of Family Income and Expenditure 1999,1994,1989, and 1984.
#: We are unable to obtain interior value for turning age in this specification. 31 is the minimum possible value in our estimation.
*, **, and *** represent significance at 10%, 5%, and 1% level, respectively. t value in parenthesis.
Under all specifications, we drop observations to households that: (1) have a female head of household;
(2) are single households; (3) work in the agricultural sector; (4) are self-employed; and (5) are company directors or firm managers
Spec1 uses the variances of the raw data.
Spec2 uses the variances of the residuals after controlling for the number of household members.
Spec3: Controlled for the number of household members and the area information.
Spec4: Controlled as in Spec3 and employment status of the family head.
Spec5: Controlled as in Spec4 and the number of employed family members.
Spec6: Controlled as in Spec5 and industry the household head is working in
Spec7: Controlled as in Spec5 and household head's type of job
Spec8: Controlled as in Spec7 and the industry the head is working for.

Table 2
Estimation of Permanent Income Shocks with Auto Correlation

	Spec1	Spec2	Spec3	Spec4	Spec5	Spec6	Spec7	Spec8
AR1 Coefficient	1.049***	1.0383***	1.0351***	1.0247***	1.0166***	1.0227***	1.0128***	1.0152***
	(171.10)	(186.65)	(181.57)	(171.34)	(156.07)	(160.71)	(136.37)	(137.20)
Variances of Permanent Shock	0.00056412	0.00037096	0.00069714	0.00081505	0.0013067	0.00070135	0.00052133	0.0002648
	(0.56)	(0.34)	(0.65)	(0.90)	(1.07)	(0.64)	(0.49)	(0.25)
89--94 year dummy	0.0012108	0.0020887	0.00092223	0.0012772	0.0016857	0.0013456	0.0013905	0.0016234
	(0.95)	(1.64)	(0.79)	(1.18)	(1.47)	(1.25)	(1.24)	(1.49)
94--99 year dummy	0.00013679	-0.00012631	-0.00020652	0.00081543	0.0007063	0.00098472	0.0022628**	0.0023226**
	(0.11)	(-0.99)	(-0.18)	(1.06)	(0.61)	(0.92)	(2.01)	(2.12)
σ	0.034716	0.031614	0.028338	0.024676	0.024934	0.024174	0.023661	0.023409
log-likelihood	180.57	189.28	199.45	212.32	211.35	214.23	216.23	217.22
Fixed Effects & Transitivity Shocks	0.130568	0.119139375	0.1014679	0.100830625	0.0838144	0.07842955	0.07807105	0.0732457

Data : National Survey of Family Income and Expenditure 1999,1994,1989, and 1984.
#: We are unable to obtain interior value for turning age in this specification. 31 is the minimum possible value in our estimation.
*, **, and *** represent significance at 10%, 5%, and 1% level, respectively. t value in parenthesis.
Under all specifications, we drop observations to households that: (1) have a female head of household;
(2) are single households; (3) work in the agricultural sector; (4) are self-employed; and (5) are company directors or firm managers
Spec1 uses the variances of the raw data.
Spec2 uses the variances of the residuals after controlling for the number of household members.
Spec3: Controlled for the number of household members and the area information.
Spec4: Controlled as in Spec3 and employment status of the family head.
Spec5: Controlled as in Spec4 and the number of employed family members.
Spec6: Controlled as in Spec5 and industry the household head is working in
Spec7: Controlled as in Spec5 and household head's type of job
Spec8: Controlled as in Spec7 and the industry the head is working for.

Table 3 Vuong's Nonnested Tests

	Vuong's Statistics	P-Values	Chosen Model
Spec1	-3.16	0.0008	Model 1
Spec2	-2.68	0.0037	Model 1
Spec3	-2.76	0.0029	Model 1
Spec4	-2.14	0.016	Model 1
Spec5	-2.12	0.017	Model 1
Spec6	-2.67	0.0038	Model 1
Spec7	-2.97	0.0015	Model 1
Spec8	-3.61	0.00016	Model 1

Vuong's Statistics show the value for AR1 Model versus Nonlinear Model.
Spec1 uses the variances of the raw data.
Spec2 uses the variances of the residuals after controlling for the number of household members.
Spec3: Controlled for the number of household members and the area information.
Spec4: Controlled as in Spec3 and employment status of the family head.
Spec5: Controlled as in Spec4 and the number of employed family members.
Spec6: Controlled as in Spec5 and industry the household head is working in
Spec7: Controlled as in Spec5 and household head’s type of job
Spec8: Controlled as in Spec7 and the industry the head is working for.

Table 4
Estimation of Permanent Income Shocks without Year Effects

	Spec1	Spec2	Spec3	Spec4	Spec5	Spec6	Spec7	Spec8
Variances of Permanent Shocks	0.0014483***	0.0018296***	0.002027***	0.0020211***	0.0025257***	0.0022108***	0.0013699***	0.0014516***
	(9.80)	(13.26)	(5.30)	(4.82)	(5.86)	(5.78)	(4.04)	(12.15)
Variance Growth Rate after Turning Age	0.20003***	0.20993***	0.19170***	0.16612***	0.19546***	0.21894***	0.30676***	0.29797***
	(58.24)	(110.43)	(9.19)	(7.10)	(7.29)	(8.21)	(7.11)	(312.77)
Turning Age	45	48	48	48	51	51	52	52
σ	0.025331	0.024551	0.021466	0.020015	0.020335	0.018628	0.0181	0.017965
log-likelihood	209.88	212.79	225.28	231.79	230.31	238.47	241.14	241.84
Fixed Effects & Transitivity Shocks	0.130568	0.119139375	0.1014679	0.100830625	0.0838144	0.07842955	0.07807105	0.0732457

Data : National Survey of Family Income and Expenditure 1999,1994,1989, and 1984.
#: We are unable to obtain interior value for turning age in this specification. 31 is the minimum possible value in our estimation.
*, **, and *** represent significance at 10%, 5%, and 1% level, respectively. t value in parenthesis.
Under all specifications, we drop observations to households that: (1) have a female head of household;
(2) are single households; (3) work in the agricultural sector; (4) are self-employed; and (5) are company directors or firm managers
Spec1 uses the variances of the raw data.
Spec2 uses the variances of the residuals after controlling for the number of household members.
Spec3: Controlled for the number of household members and the area information.
Spec4: Controlled as in Spec3 and employment status of the family head.
Spec5: Controlled as in Spec4 and the number of employed family members.
Spec6: Controlled as in Spec5 and industry the household head is working in
Spec7: Controlled as in Spec5 and household head's type of job
Spec8: Controlled as in Spec7 and the industry the head is working for.

Table 5: Normalized Financial Wealth at 25 Years Old

Year	99	94	89	84
Average Financial Wealth at 25 Years Old	0.4726	0.4176	0.4480	0.4786
Variance of Financial Wealth at 25 Years Old	0.5937	0.4154	0.3284	0.3170

Data : National Survey of Family Income and Expenditure 1999,1994,1989, and 1984

Table 6

Method of Simulated Moments Estimation using the Variance Level (4 years pooled)

Interest Rate:	3%				4%				5%			
Fraction of Fixed Effect	estimated	0.4	0.5	0.6	estimated	0.4	0.5	0.6	estimated	0.4	0.5	0.6
Estimated γ	2.8326	0.4533	1.1496	2.7863	2.0911	0.2940	0.7690	2.1291	1.0891	0.1261	0.2929	1.2091
	(2.33)	(12.12)	(6.66)	(6.09)	(2.04)	(11.47)	(6.11)	(5.05)	(1.78)	(9.87)	(6.10)	(3.50)
Estimated χ	0.0170	0.0181	0.0182	0.0170	0.0170	0.0179	0.0180	0.0170	0.0169	0.0176	0.0177	0.0169
	(19.07)	(16.10)	(17.80)	(22.45)	(18.58)	(15.94)	(17.79)	(22.67)	(17.75)	(15.59)	(17.23)	(23.04)
Log Likelihood	382.06	373.39	372.58	382.05	382.07	375.15	373.82	382.06	383.11	377.25	376.93	383.00
Estimated Fraction of Fixed Effect	0.6023	-	-	-	0.5977	-	-	-	0.5908	-	-	
Std. of Fixed Effect	0.0320	-	-	-	0.0305	-	-	-	0.0247	-	-	-

Note : t value in parenthesis.

: Fixed effects plus transitory shocks are 0.10541(99), 0.10193(94), 0.10923(89) and 0.10617(84).

: On the left column for each interest rate, the fraction of fixed effect is estimated by the maximum likellihood.

Table 7

Method of Simulated Moments Estimation using the Variance Level (3 years pooled)

Interest Rate:	3%				4%				5%			
Fraction of Fixed Effect	estimated	0.4	0.5	0.6	estimated	0.4	0.5	0.6	estimated	0.4	0.5	0.6
Estimated γ	2.8344	0.4984	1.3906	3.0663	2.0930	0.3215	0.9493	2.3824	1.0883	0.1360	0.3981	1.4253
	(2.23)	(10.62)	(5.80)	(6.12)	(1.93)	(10.26)	(5.25)	(5.05)	(1.68)	(9.43)	(4.76)	(3.44)
Estimated χ	0.0161	0.0175	0.0172	0.0162	0.0162	0.0173	0.0171	0.0162	0.0161	0.0168	0.0167	0.0161
	(16.49)	(13.44)	(15.70)	(20.04)	(16.01)	(13.21)	(15.63)	(20.10)	(15.15)	(12.66)	(14.85)	(20.00)
Log Likelihood	292.39	283.56	285.62	292.22	292.29	285.17	286.46	291.97	292.98	288.14	288.53	292.36
Estimated Fraction of Fixed Effect	0.5881	-	-	-	0.5836	-	-	-	0.5766	-	-	-
Std. of Fixed Effect	0.0351	-	-	-	0.0338	-	-	-	0.0279	-	-	-

Note : t value in parenthesis.

: 1999's consumption variances are excluded.

: Fixed effects plus transitory shocks are 0.10193(94), 0.10923(89) and 0.10617(84).

: On the left column for each interest rate, the fraction of fixed effect is estimated by the maximum likellihood.

Table 8a: Method of Simulated Moments (1999)

Interest Rate:	3%			4%			5%		
Fraction of Fixed Effect	0.4	0.5	0.6	0.4	0.5	0.6	0.4	0.5	0.6
Estimated γ	0.3606	0.6134	2.0220	0.2329	0.4003	1.4576	0.1013	0.1717	0.7201
	(5.64)	(3.87)	(1.96)	(5.37)	(3.59)	(1.63)	(4.03)	(2.99)	(1.12)
Estimated χ	0.0191	0.0198	0.0191	0.0190	0.0196	0.0190	-0.0195	0.0192	0.0188
	(8.87)	(9.13)	(9.56)	(8.86)	(9.15)	(9.72)	(-8.22)	(9.13)	(10.58)
Log Likelihood	91.419	90.067	91.468	91.539	90.410	91.623	90.687	91.145	92.063

Note : t value in parenthesis.

: Fixed effect plus transitory shock is 0.10541.

Table 8b: Method of Simulated Moments (1994)

Interest Rate:	3%			4%			5%		
Fraction of Fixed Effect	0.4	0.5	0.6	0.4	0.5	0.6	0.4	0.5	0.6
Estimated γ	0.3876	0.7910	2.0133	0.2508	0.5246	1.4469	0.1095	0.2172	0.7261
	(7.73)	(5.02)	(4.28)	(7.32)	(4.73)	(3.75)	(5.76)	(4.51)	(2.94)
Estimated χ	0.0144	0.0143	0.0128	0.0143	0.0142	0.0127	0.0148	0.0140	0.0125
	(6.06)	(6.84)	(10.15)	(6.06)	(6.74)	(10.34)	(6.92)	(6.45)	(10.26)
Log Likelihood	101.618	101.831	105.924	101.859	102.190	106.099	100.628	102.619	106.663

Note : t value in parenthesis.

: Fixed effect plus transitory shock is 0.10193.

Table 8c: Method of Simulated Moments (1989)

Interest Rate:	3%			4%			5%		
Fraction of Fixed Effect	0.4	0.5	0.6	0.4	0.5	0.6	0.4	0.5	0.6
Estimated γ	0.4947	1.3873	3.2930	0.3230	0.9582	2.6218	0.1418	0.4241	1.6524
	(6.54)	(3.31)	(3.36)	(6.13)	(3.00)	(2.73)	(5.12)	(2.72)	(1.82)
Estimated χ	0.0166	0.0164	0.0149	0.0164	0.0162	0.0148	0.0165	0.0160	0.0147
	(7.32)	(9.13)	(11.01)	(7.41)	(9.09)	(11.12)	(8.15)	(8.72)	(11.44)
Log Likelihood	96.455	96.900	100.403	96.798	97.260	100.509	96.687	97.790	100.760

Note : t value in parenthesis.

: Fixed effect plus transitory shock is 0.10923.

Table 8d: Method of Simulated Moments (1984)

Interest Rate:	3%			4%			5%		
Fraction of Fixed Effect	0.4	0.5	0.6	0.4	0.5	0.6	0.4	0.5	0.6
Estimated γ	0.7016	2.2865	4.1604	0.4403	1.6706	3.4609	0.1663	0.8102	2.3440
	(4.26)	(2.70)	(3.11)	(4.29)	(2.30)	(2.50)	(4.69)	(1.73)	(1.70)
Estimated χ	0.0201	0.0188	0.0192	0.0197	0.0188	0.0194	0.0184	0.0186	0.0195
	(8.24)	(9.65)	(11.17)	(7.94)	(9.49)	(11.01)	(6.81)	(8.79)	(10.65)
Log Likelihood	89.633	91.896	91.193	90.305	91.890	90.852	92.849	92.271	90.652

Note : t value in parenthesis.

: Fixed effect plus transitory shock is 0.10617.

Table 9a: Method of Simulated Moments Estimation using Slopes of the Variances (4 years pooled)

Interest Rate:	3%			4%			5%		
Fraction of Fixed Effect	0.4	0.5	0.6	0.4	0.5	0.6	0.4	0.5	0.6
Estimated γ	4.2328	4.4449	5.0712	3.5284	3.6887	4.1139	2.4821	2.5632	2.8636
	(2.69)	(2.48)	(2.33)	(2.28)	(2.07)	(1.93)	(1.74)	(1.57)	(1.43)
Estimated χ	0.0171	0.0180	0.0190	0.0171	0.0179	0.0189	0.0170	0.0179	0.0188
	(17.56)	(18.31)	(17.12)	(17.51)	(18.57)	(18.30)	(17.41)	(19.02)	(19.70)
Log Likelihood	246.27	241.66	236.74	246.55	241.93	236.99	246.99	242.41	237.45

Note : t value in parenthesis.

: Fixed effects plus transitory shocks are 0.10541(99), 0.10193(94), 0.10923(89) and 0.10617(84).

Table 9b: Method of Simulated Moments Estimation using Slopes of the Variances (3 years pooled)

Interest Rate:	3%			4%			5%		
Fraction of Fixed Effect	0.4	0.5	0.6	0.4	0.5	0.6	0.4	0.5	0.6
Estimated γ	4.6099	4.8055	5.5114	3.9511	4.0589	4.4145	2.9935	3.0294	3.2006
	(2.28)	(2.09)	(2.02)	(1.91)	(1.78)	(1.66)	(1.46)	(1.34)	(1.24)
Estimated χ	0.0177	0.0185	0.0194	0.0176	0.0184	0.0193	0.0175	0.0183	0.0192
	(13.52)	(14.16)	(13.53)	(13.40)	(14.30)	(14.46)	(13.29)	(14.37)	(15.22)
Log Likelihood	162.19	159.42	156.43	162.48	159.69	156.67	162.80	160.03	157.01

Note : t value in parenthesis.

: 1999's consumption variances are excluded.

: Fixed effects plus transitory shocks are 0.10193(94), 0.10923(89) and 0.10617(84).

Table 10: Method of Simulated Moments Estimation using the Variance Level (Linear Regression, 4 Years Pooled)

Interest Rate:	4%		
Fraction of Fixed Effect	0.4	0.5	0.6
Estimated γ	2.8585	10.4537	24.5987
	(3.00)	(4.80)	(4.51)
Estimated χ	-0.0271	-0.0270	0.0246
	(-12.34)	(-15.05)	(15.04)
Log Likelihood	315.18	315.85	329.14
Std. of Fixed Effect	-	-	-
Vuong's Statistics	3.54	3.37	3.25
Chosen Model	Nonlinear	Nonlinear	Nonlinear

Note : t value in parenthesis.

: Fixed effects plus transitory shocks are 0.06743(99), 0.07085(94), 0.07594(89) and 0.07925(84).

Appendix Table 1

Explanation of Control Variables for Each Specification

Spec1	Constant
Spec2	Spec1 + the Number of Household Member Dummies
Spec3	Spec2 + Area Information
Spec4	Spec3 + Employment Status Dummy of Household Head
Spec5	Spec4 + the Number of Employed Household Member Dummies
Spec6	Spec5 + Industrial Dummy of Household Head
Spec7	Spec4 + Type of Job Dummy of Household Head
Spec8	Spec6 + Spec7
Area Information	Dummy for the Major Three Metropolitan Areas: 1 or 0
	Dummies for the City Groups (the size of population):
	Five Ranked, Four dummies
	Dummies for 10 Districts: Nine Dummies
	Dummies for the Large Four Cities: Three Dummies
	Dummy for the Largest Three Cities: 1 or 0
	Dummy for the location of Prefectural Administration Center: 1 or 0
Industrial Dummies	Agricultural, Forestry, and Fisheries Workers*
Each of them takes 1 or 0	Mining
	Construction
	Manufacturing
	Electricity, Gas, Water, and Heat
	Retail, Wholesale, and Hospitality
	Financial Institutions
	Real Estate
	Service
	Public Officer
	Other (including Unemployed)
Job Type Dummies	Full Time Nonoffice Workers
Each of them takes 1 or 0	Part Time Nonoffice Workers
	Office Workers
	Individual Proprietors*
	Corporate Administrators*
	Agricultural, Forestry, and Fisheries Workers*
	Professional Service
	Other Occupation
	No Occupation

See The NSFIE Reports published by Statistical Bureau, Ministry of Public Management, Home Affairs, Posts and Telecommunications, Japan for more detailed explanation for each variable. Variables with * are not included in our regressions since we exclude households that correspond with the categories.

Appendix Table 2
Variances of Income and Consumption

	Variances of household income (log)				Variances of non durable consumption (log)			
Age	1999	1994	1989	1984	1999	1994	1989	1984
25	0.1042	0.1048	0.1050	0.0919	0.0964	0.1095	0.0960	0.0776
26	0.0939	0.0935	0.1042	0.1239	0.0929	0.1122	0.1063	0.0931
27	0.1034	0.1238	0.1122	0.1127	0.0836	0.0916	0.0873	0.0987
28	0.1106	0.0908	0.1091	0.0966	0.1206	0.1075	0.1217	0.1014
29	0.1150	0.0967	0.1157	0.1058	0.1161	0.1010	0.1116	0.0940
30	0.1228	0.1099	0.1175	0.1297	0.1251	0.0999	0.0974	0.0796
31	0.1223	0.1232	0.1042	0.1182	0.1064	0.1061	0.1109	0.0856
32	0.1206	0.1130	0.1144	0.1170	0.1179	0.1060	0.1015	0.0838
33	0.1005	0.1118	0.1078	0.1251	0.1166	0.1001	0.0986	0.0955
34	0.1087	0.1246	0.1428	0.1156	0.1187	0.1006	0.0945	0.0925
35	0.1200	0.1159	0.1215	0.1193	0.1078	0.0994	0.0997	0.0792
36	0.1085	0.1227	0.1253	0.1192	0.1073	0.1034	0.0927	0.0850
37	0.1347	0.1273	0.1381	0.1228	0.1113	0.0896	0.0955	0.0791
38	0.1209	0.1298	0.1231	0.1344	0.1095	0.0968	0.0914	0.0784
39	0.1149	0.1407	0.1388	0.1357	0.1018	0.0992	0.0876	0.0777
40	0.1376	0.1354	0.1316	0.1107	0.0964	0.1012	0.0902	0.0759
41	0.1242	0.1420	0.1415	0.1652	0.0946	0.0894	0.0936	0.0828
42	0.1308	0.1596	0.1591	0.1591	0.0894	0.0961	0.0943	0.0839
43	0.1268	0.1397	0.1399	0.1598	0.1135	0.0930	0.1021	0.0852
44	0.1491	0.1532	0.1369	0.1705	0.1061	0.0955	0.1033	0.0816
45	0.1451	0.1865	0.1655	0.1741	0.1001	0.0966	0.1018	0.0885
46	0.1477	0.1530	0.1538	0.1782	0.1089	0.1093	0.0950	0.0971
47	0.1459	0.1670	0.1552	0.1537	0.1104	0.1132	0.1068	0.0971
48	0.1515	0.1639	0.1526	0.1704	0.1097	0.1041	0.1067	0.0918
49	0.2022	0.1858	0.1582	0.1823	0.1180	0.1126	0.1131	0.1137
50	0.1958	0.1628	0.2240	0.1905	0.1171	0.1108	0.1241	0.1149
51	0.1709	0.1616	0.1715	0.1922	0.1316	0.1228	0.1302	0.1404
52	0.1655	0.2342	0.1949	0.1939	0.1404	0.1322	0.1597	0.1281
53	0.1945	0.1788	0.2039	0.1947	0.1531	0.1350	0.1466	0.1424
54	0.1832	0.2048	0.2073	0.1994	0.1605	0.1443	0.1494	0.1433
55	0.1926	0.2037	0.1999	0.2222	0.1560	0.1520	0.1420	0.1846
56	0.2146	0.2380	0.2472	0.2087	0.1444	0.1748	0.1630	0.1694
57	0.2469	0.2643	0.2829	0.2441	0.1800	0.1592	0.1704	0.1496
58	0.2907	0.2854	0.2629	0.3020	0.1666	0.1843	0.1748	0.2071
59	0.3018	0.3398	0.2828	0.2798	0.1897	0.1876	0.1530	0.1742
60	0.3415	0.2549	0.2691	0.2477	0.1921	0.1746	0.1799	0.1766
61	0.2973	0.2650	0.2471	0.2310	0.1953	0.1671	0.1677	0.1874
62	0.2998	0.2650	0.2636	0.3053	0.1974	0.2053	0.1812	0.1970
63	0.2721	0.2734	0.2527	0.2544	0.2152	0.1775	0.1733	0.1838
64	0.2732	0.2990	0.2416	0.2789	0.1781	0.1747	0.1930	0.1684
65	0.2796	0.2506	0.2297	0.2255	0.1888	0.1613	0.1872	0.1660
66	0.2542	0.2300	0.2731	0.2580	0.1859	0.1752	0.1564	0.1250
67	0.2473	0.2356	0.2866	0.2551	0.1678	0.1732	0.2023	0.1679
68	0.2249	0.2146	0.2533	0.3003	0.1828	0.1785	0.1625	0.1704
69	0.2120	0.2819	0.2668	0.2226	0.1656	0.1742	0.1901	0.1467
70	0.2233	0.2220	0.2428	0.2759	0.1657	0.1642	0.2314	0.1456
71	0.1945	0.2101	0.2653	0.2669	0.1667	0.1767	0.1867	0.2105
72	0.2006	0.2219	0.3166	0.2776	0.1738	0.1769	0.1685	0.1735
73	0.2613	0.2634	0.2309	0.2732	0.1639	0.1540	0.1827	0.1589
74	0.2061	0.2266	0.2467	0.2864	0.1779	0.2135	0.1744	0.1754
75	0.2172	0.2148	0.3028	0.2139	0.2247	0.1624	0.1144	0.1532
76	0.2204	0.2079	0.2718	0.2095	0.2097	0.2051	0.1996	0.1045
77	0.1726	0.2980	0.3962	0.2959	0.1492	0.1660	0.1687	0.1393
78	0.1772	0.1852	0.2581	0.3528	0.1396	0.1985	0.2096	0.1162
79	0.1824	0.2287	0.2798	0.1736	0.1319	0.1383	0.1974	0.0873
80	0.2220	0.3422	0.3063	0.2492	0.2524	0.1981	0.1613	0.1221

Data : National Survey of Family Income and Expenditure 1999,1994,1989, and 1984.

We drop observations to households that: (1) have a female head of household; (2) are single households; (3) work in the agricultural sector; (4) are self-employed; and (5) are company directors or firm managers.

The number of children (younger than 15) is included as a control variable for consumption.

Figure 1: Age Log-Income Variance Profiles

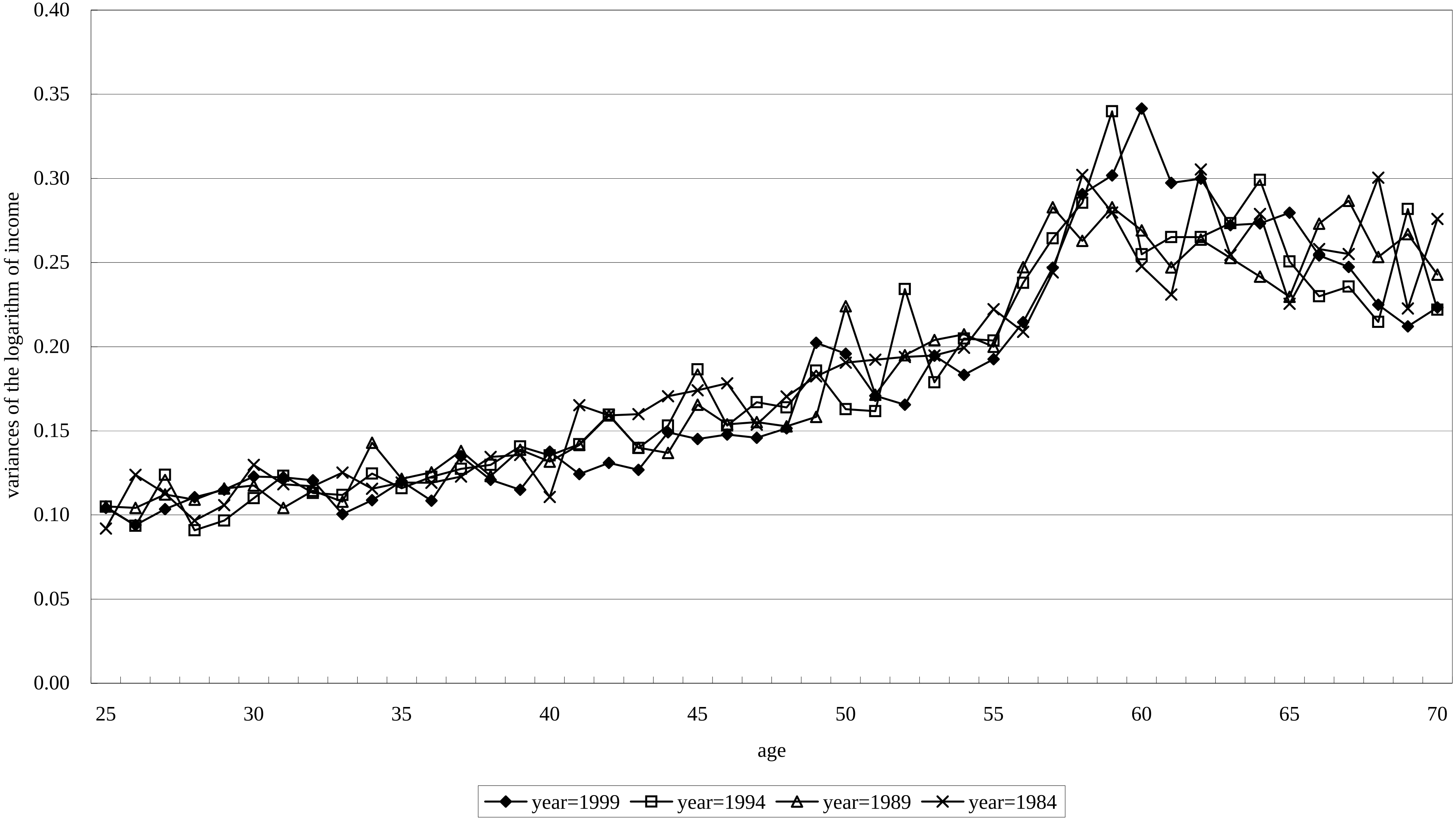

Figure2: Age Log-Consumption Variance Profile

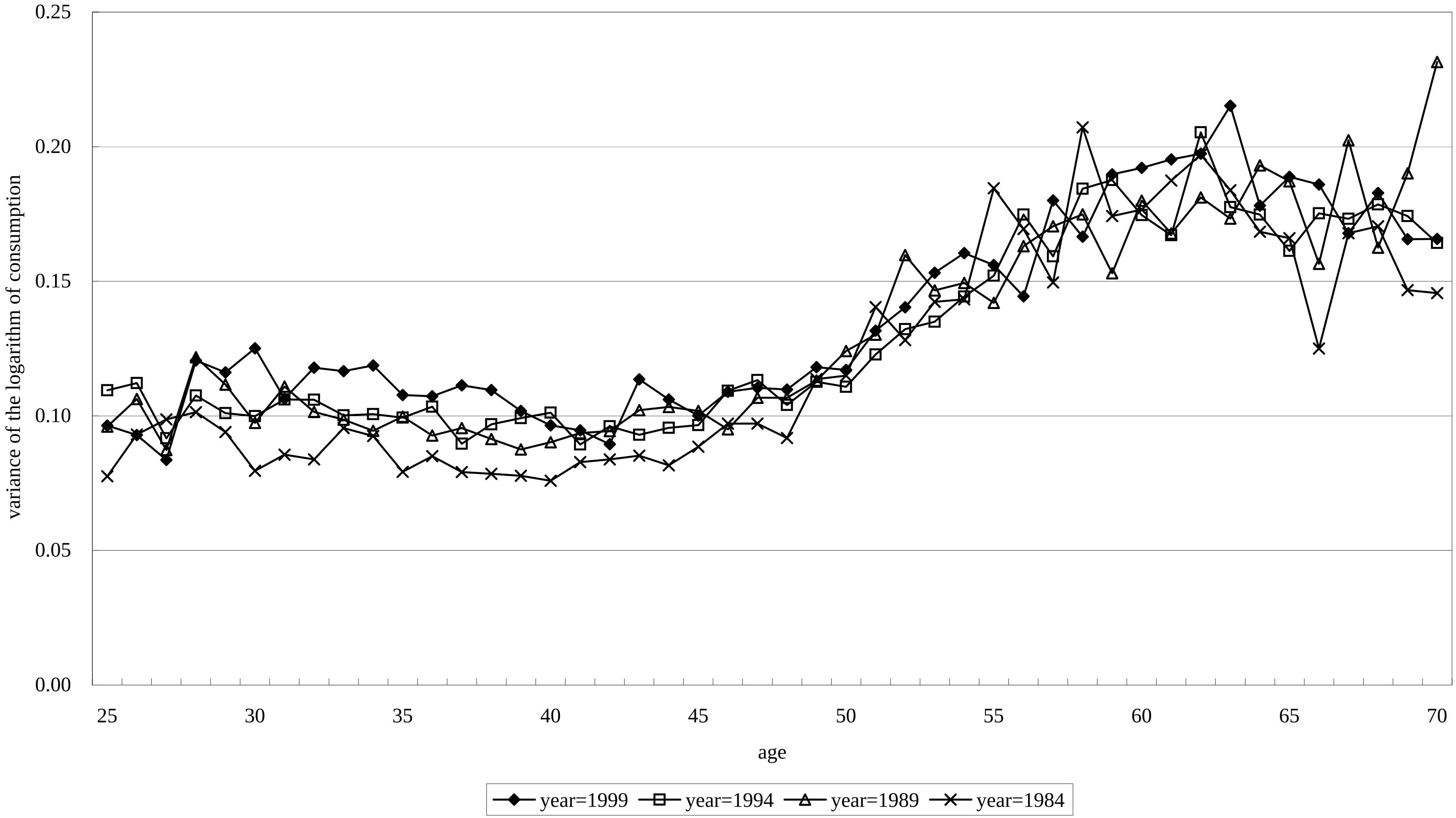

Figure 3: Consumption Function for Each Age
log utility (γ=1.0)

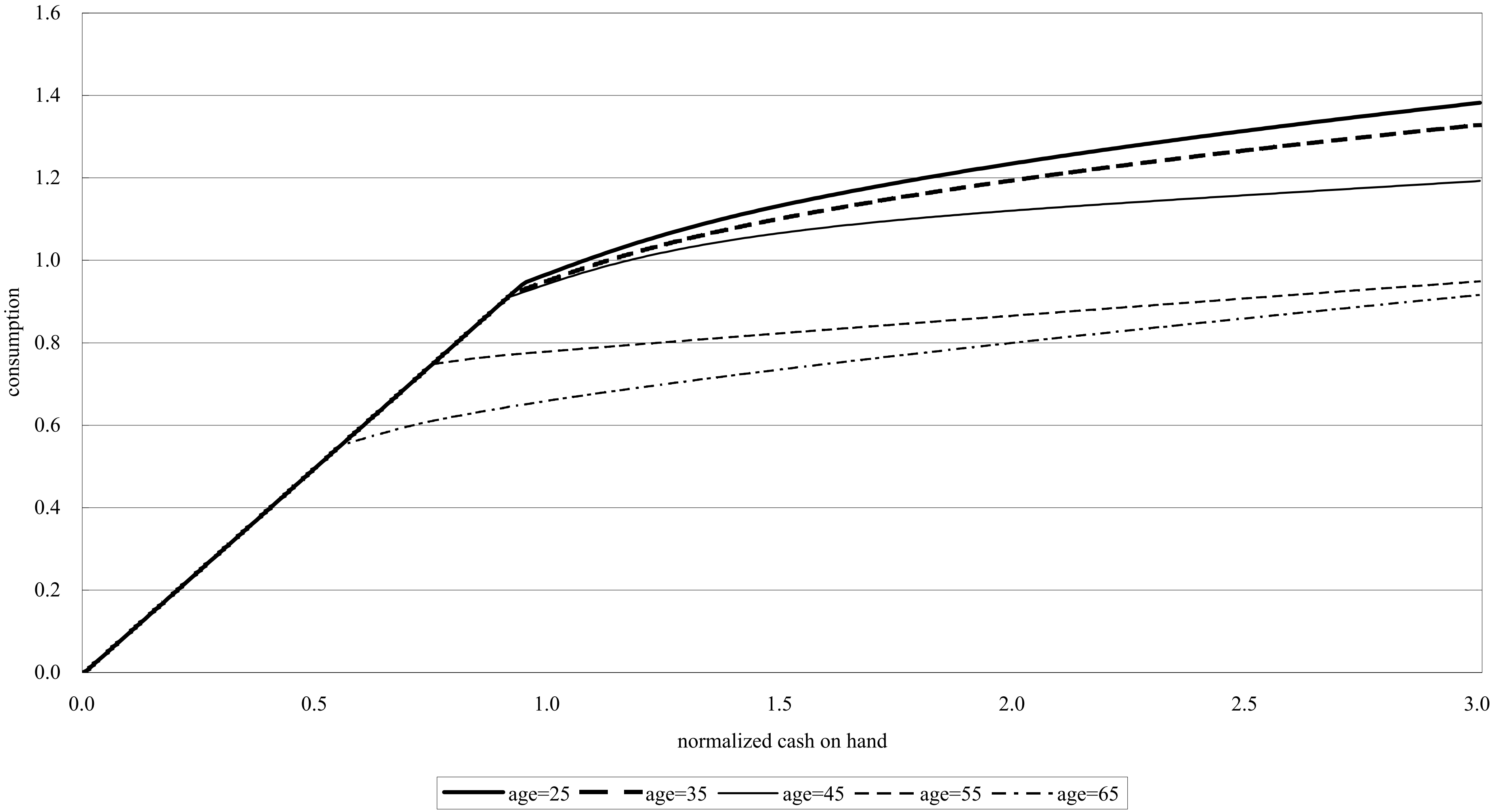

Figure 4a: Age-Income Profile in 1999 (Data and Smoothed Curve)

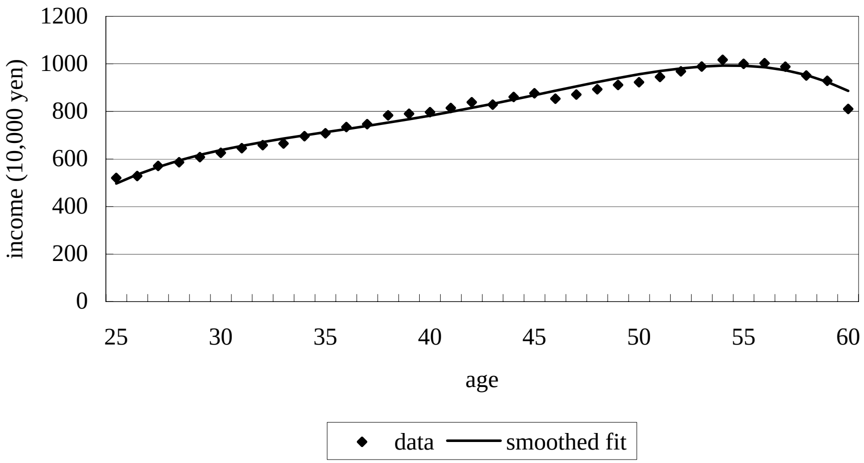

Figure 4b: Age-Income Profile in 1994 (Data and Smoothed Curve)

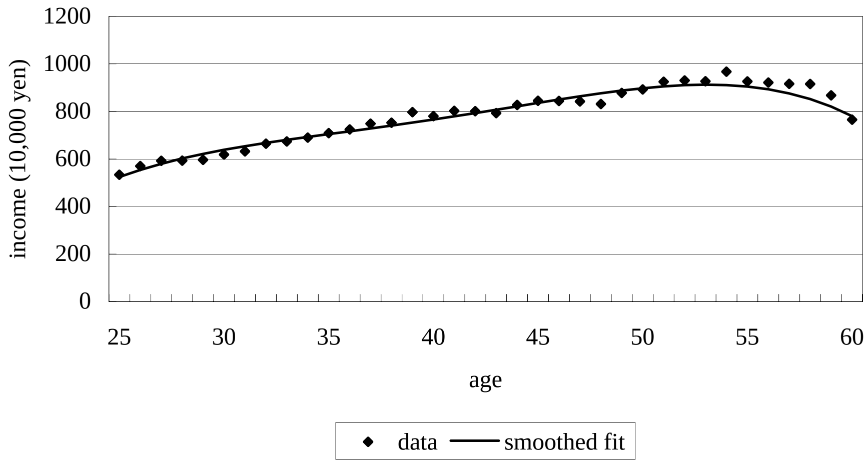

Figure 4c: Age-Income Profile in 1989 (Data and Smoothed Curve)

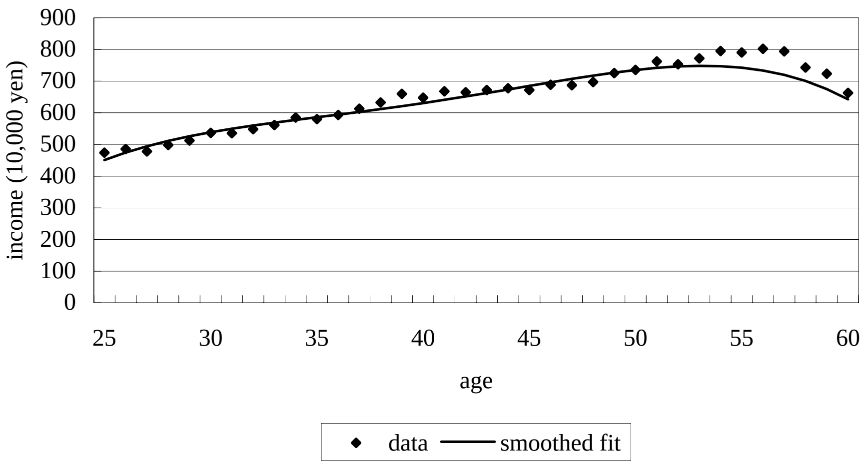

Figure 4d: Age-Income Profile in 1984 (Data and Smoothed Fit)

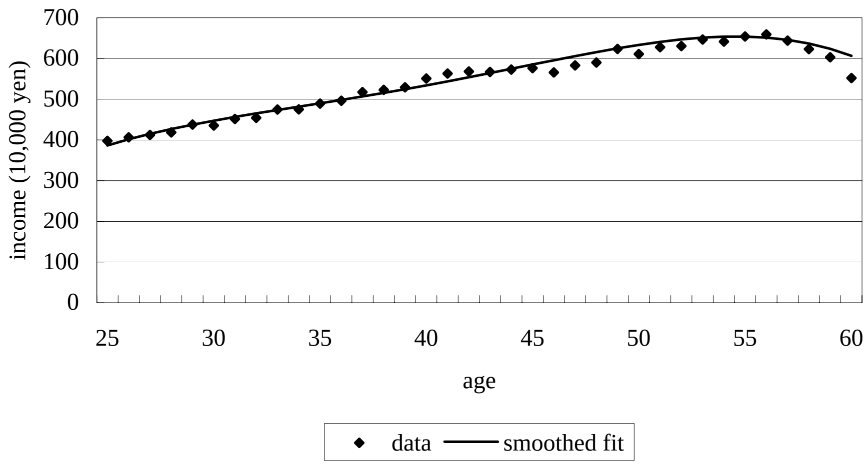

Figure 5: The Variances of the Logarithms of Consumption (Data and Simulation)

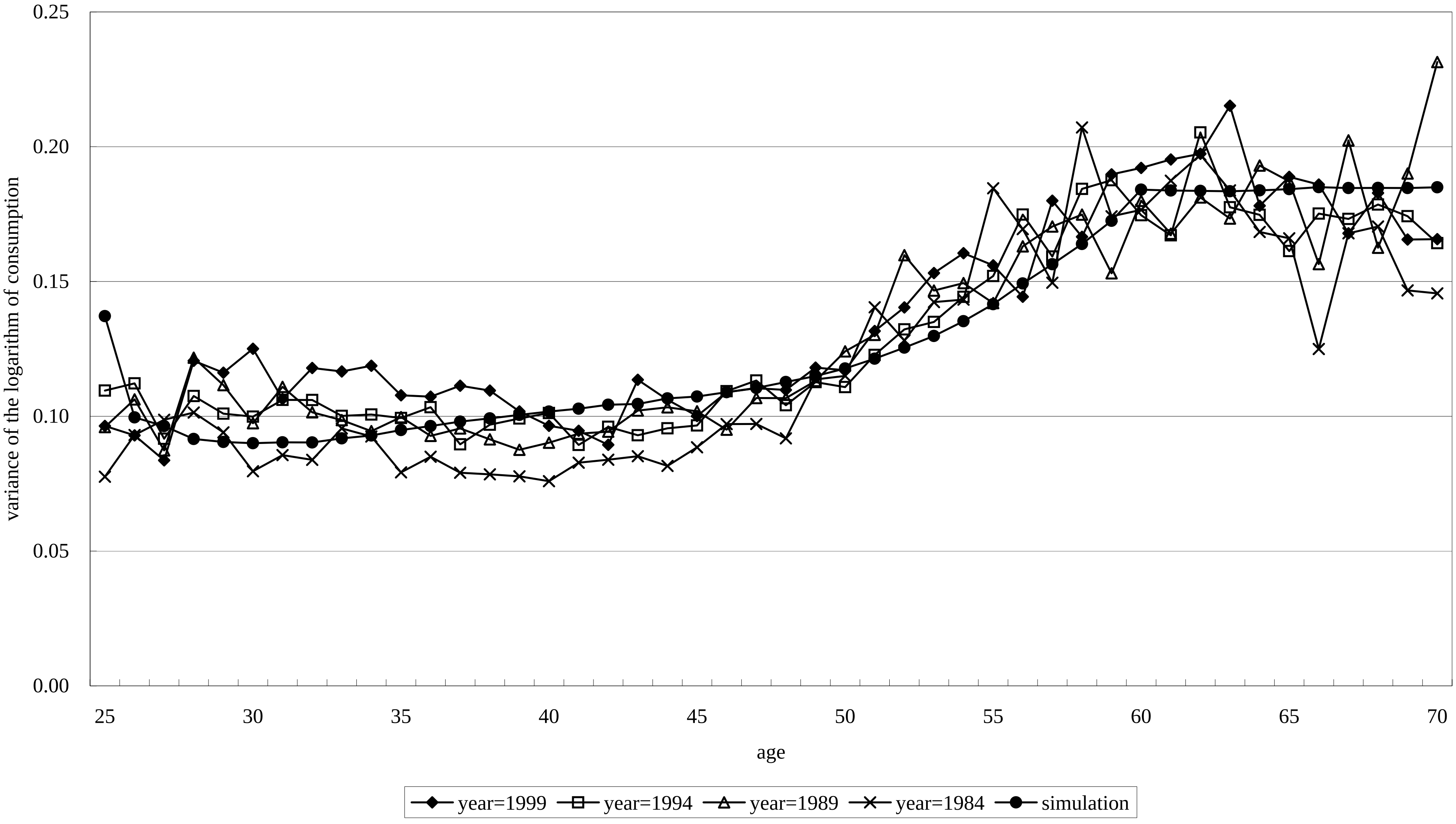

Figure 6a: Net Financial Assets over Income (Average)

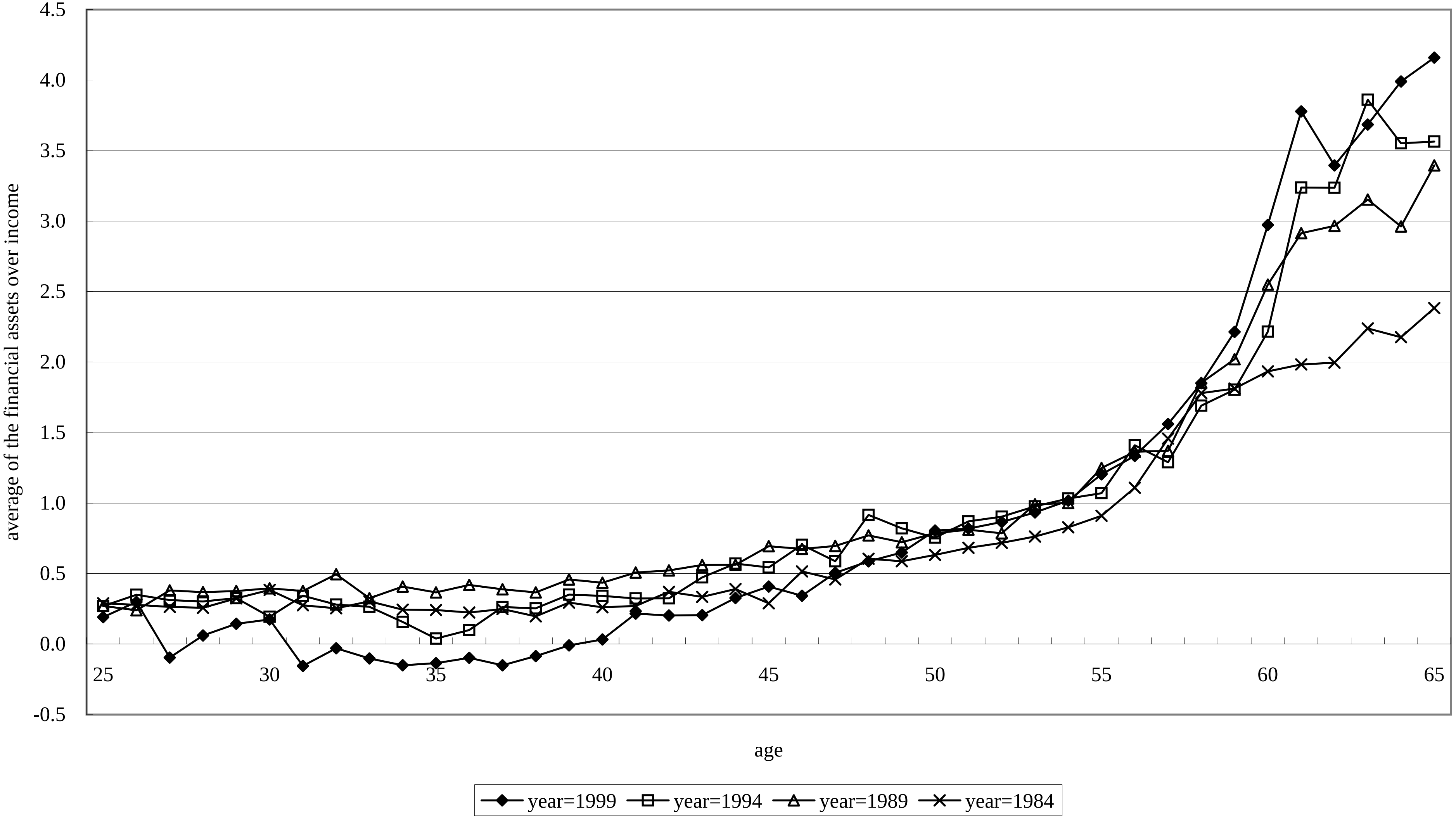

Figure 6b: Net Financial Assets over Income (Variance)

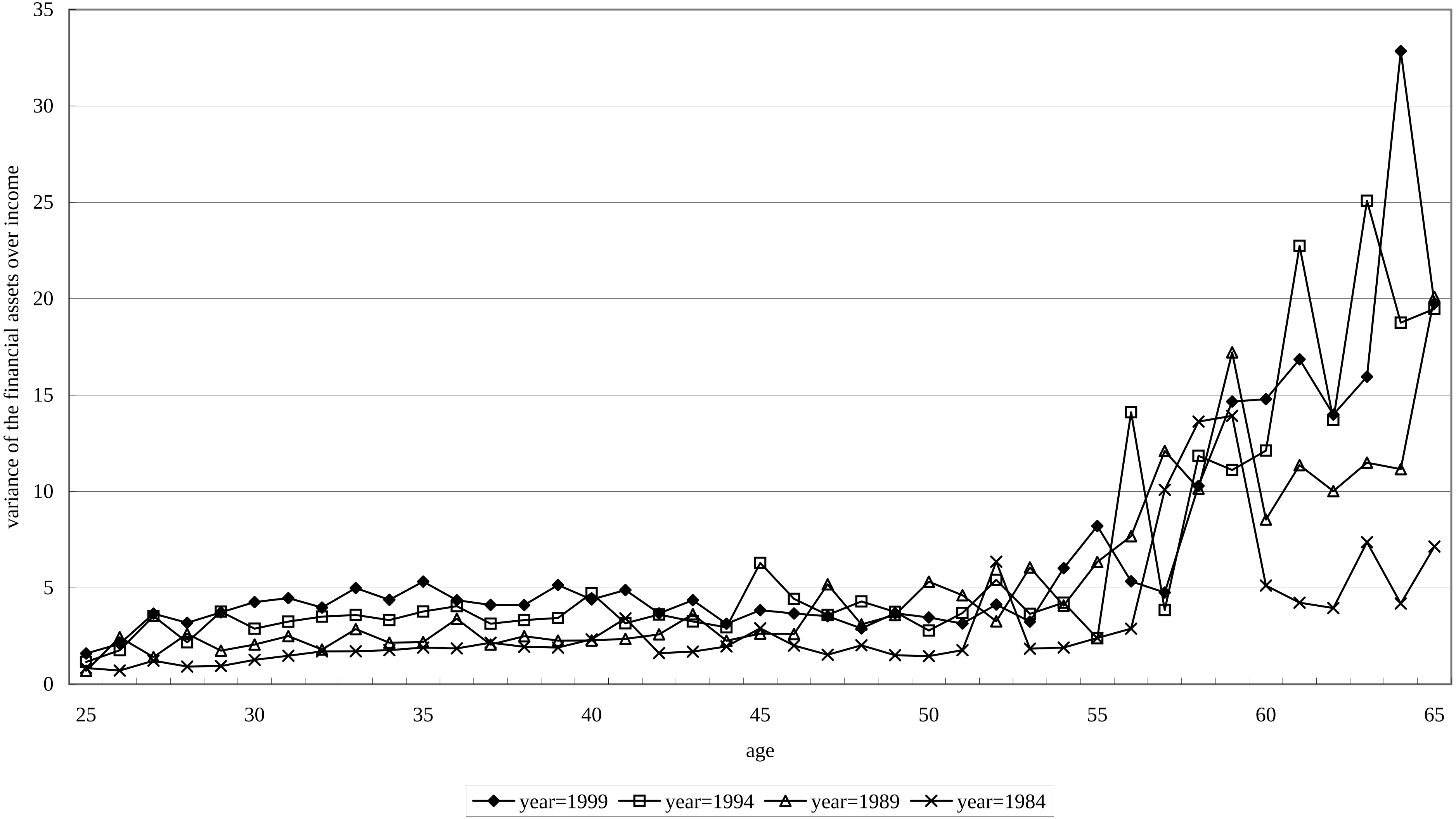

Figure 6c: Net Financial Assets Plus Housing Loan over Income (average)

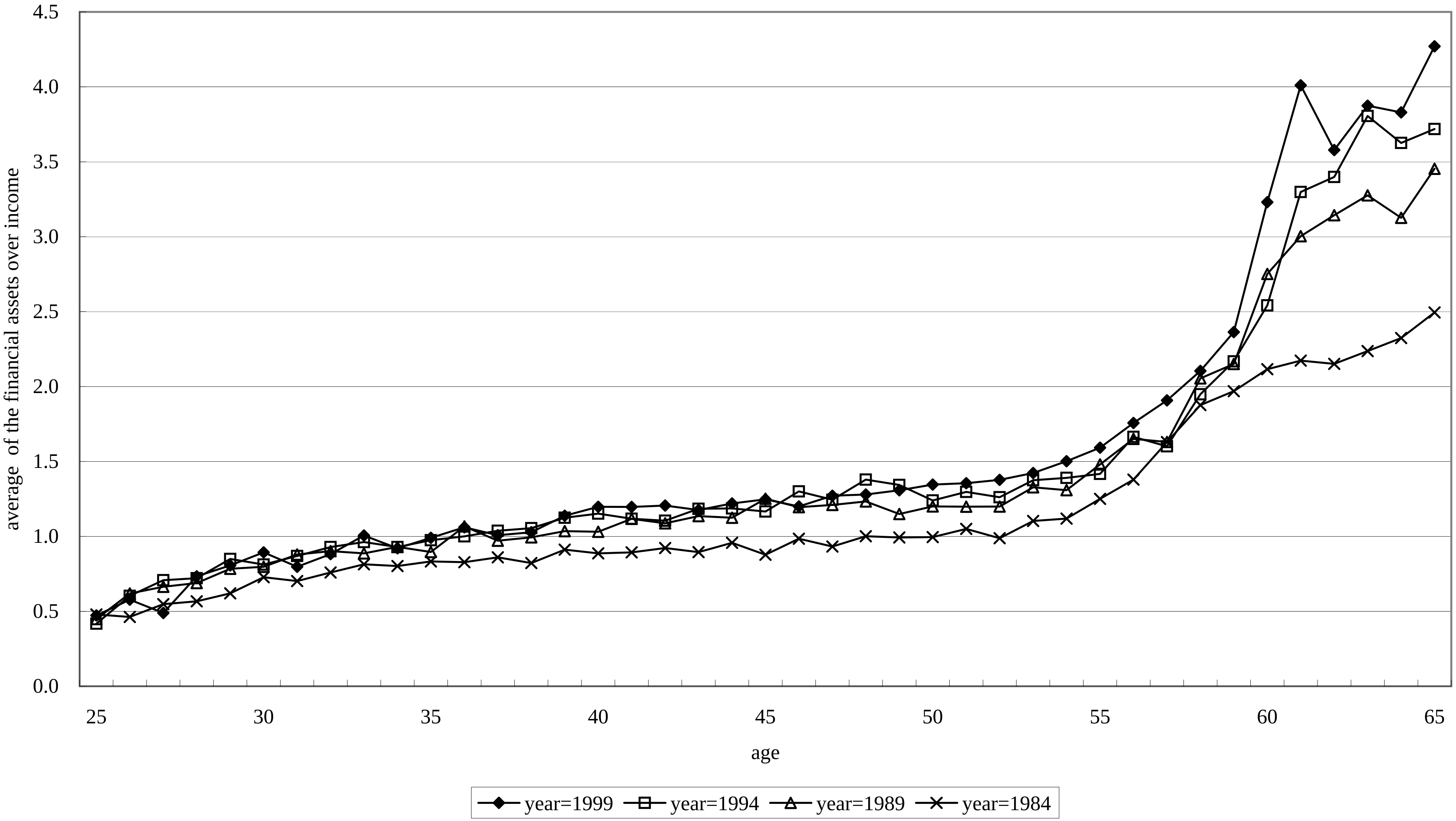

Figure 7a: The Variances of the Logarithms of Simulated Consumption for Each Cohort

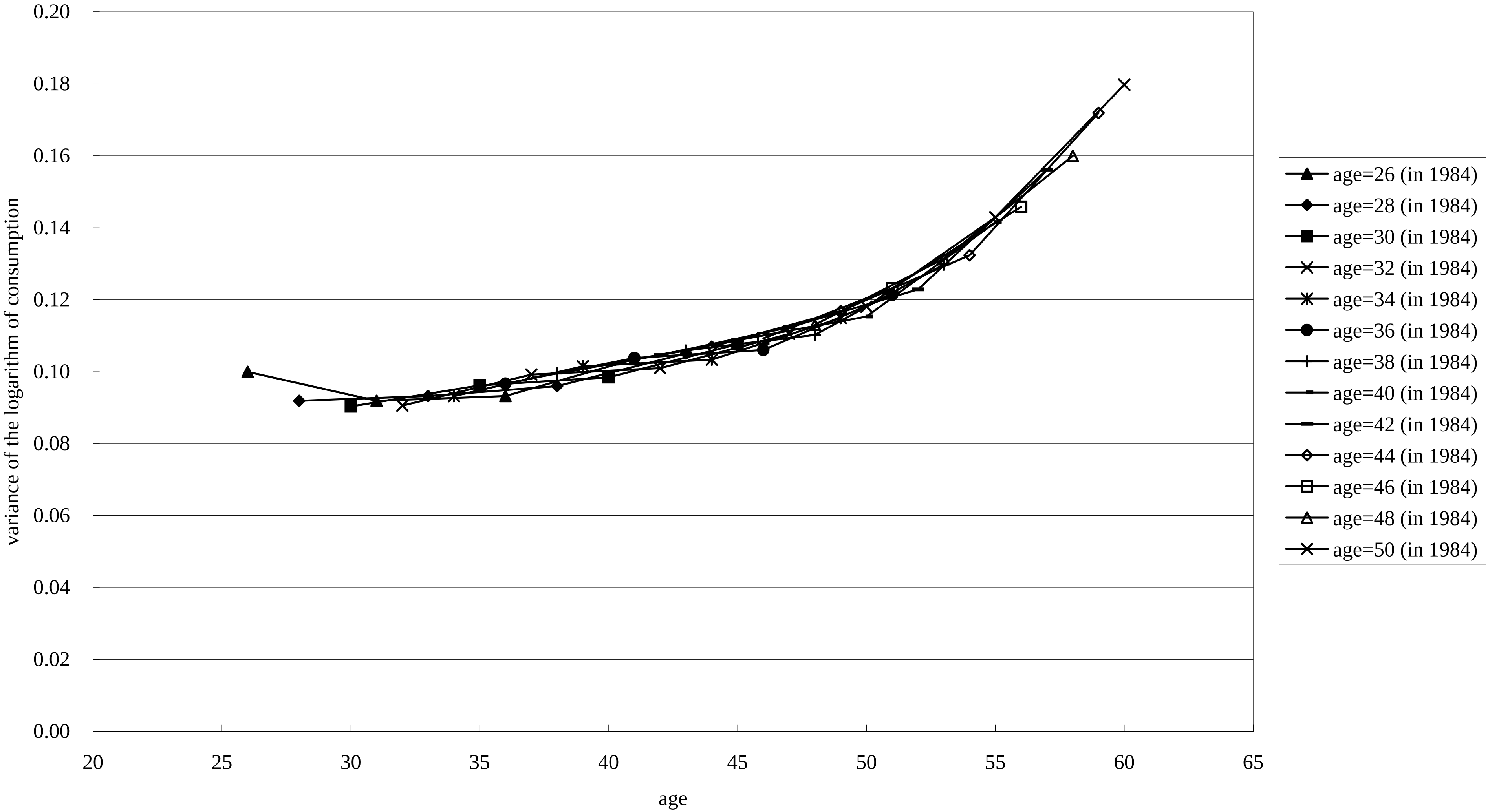

Figure 7b: The Variances of the Logarithms of Consumption for Each Cohort

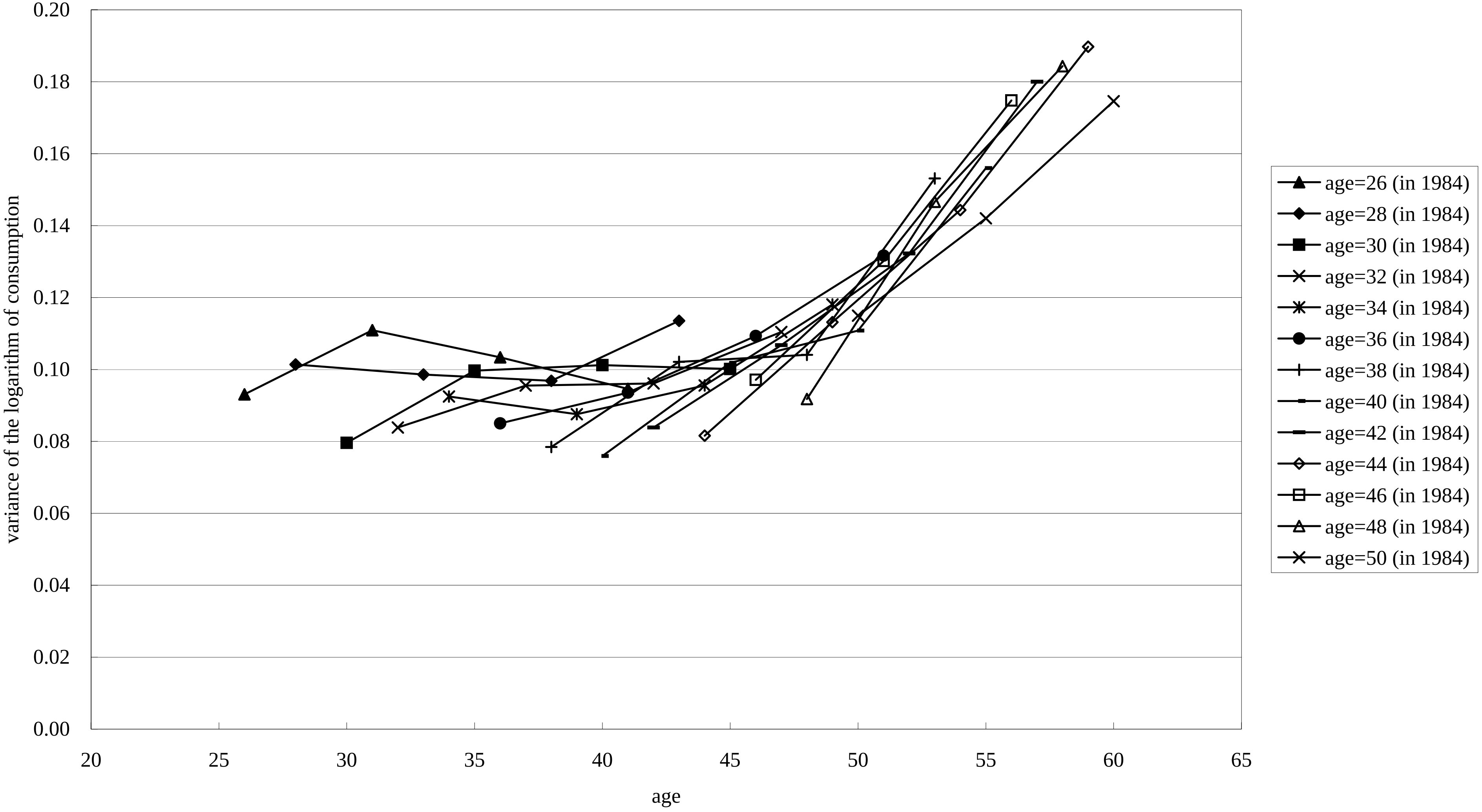

Figure 8: Linear Regression and the Variances of Consumption
(γ=10.45, Fraction of the Fixed Effect=0.5)